

involve

a journal of mathematics

Invariant measures for hybrid stochastic systems

Xavier Garcia, Jennifer Kunze, Thomas Rudelius,
Anthony Sanchez, Sijing Shao, Emily Speranza and Chad Vidden



Invariant measures for hybrid stochastic systems

Xavier Garcia, Jennifer Kunze, Thomas Rudelius,
Anthony Sanchez, Sijing Shao, Emily Speranza and Chad Vidden

(Communicated by David Royal Larson)

In this paper, we seek to understand the behavior of dynamical systems that are perturbed by a parameter that changes discretely in time. If we impose certain conditions, we can study certain embedded systems within a hybrid system as time-homogeneous Markov processes. In particular, we prove the existence of invariant measures for each embedded system and relate the invariant measures for the various systems through the flow. We calculate these invariant measures explicitly in several illustrative examples.

1. Introduction

An understanding of dynamical systems allows one to analyze the way processes evolve through time. Usually, such systems are given by differential equations that model real world phenomena. Unfortunately, these models are limited in that they cannot account for random events that may occur in application. These stochastic developments, however, may sometimes be modeled with Markov processes, and in particular with Markov chains. We can unite the two models in order to see how these dynamical systems behave with the perturbation induced by the Markov processes, creating a hybrid system consisting of the two components. Complicating matters, these hybrid systems can be described in either continuous or discrete time.

The focus of this paper is studying the way these hybrid systems behave as they evolve. We begin by defining limit sets for a dynamical system and stochastic processes. We next examine the limit sets of these hybrid systems and what happens as they approach the limit sets. Concurrently, we define invariant measures and prove their existence for hybrid systems while relating these measures to the flow. In addition, we supply examples with visuals that provide insight to the behavior of hybrid systems.

MSC2010: 34F05, 60J20, 37N20.

Keywords: dynamical systems, Markov processes, Markov chains, stochastic modeling.

Research supported by NSF grant DMS 0750986 (Kunze, Rudelius, Speranza); DMS 0750986 and DMS 0502354 (Garcia and Sanchez); and Iowa State University (Shao and Vidden).

2. The stochastic hybrid system

In this section, we define a hybrid system.

Definition 1. A Markov process X_t is called *time-homogeneous* on T if, for all $t_1, t_2, k \in T$ and for any sets $A_1, A_2 \in S$,

$$P(X_{t_1+k} \in A_1 \mid X_{t_1} \in A_2) = P(X_{t_2+k} \in A_1 \mid X_{t_2} \in A_2).$$

Otherwise, it is called *time-inhomogeneous*.

Definition 2. A Markov chain X_n is a Markov process for which perturbations occur on a discrete time set T and finite state space S .

For a Markov chain on the finite state space S with cardinality $|S|$, it is useful to describe the probabilities of transitioning from one state to another with a transition matrix

$$Q \equiv \begin{pmatrix} P_{1 \rightarrow 1} & \dots & P_{1 \rightarrow |S|} \\ \cdot & & \cdot \\ \cdot & & \cdot \\ \cdot & & \cdot \\ P_{|S| \rightarrow 1} & \dots & P_{|S| \rightarrow |S|} \end{pmatrix},$$

where $P_{i \rightarrow j}$ is the probability of transitioning from state $s_i \in S$ to state $s_j \in S$.

Also, for the purposes of this paper, we suppose that our Markov chain transitions occur regularly at times $t = nh$ for some length of time $h \in \mathbb{R}^+$ and for all $n \in \mathbb{N}$.

Definition 3. Let $\{X_n\}$, for $X_n \in S$ and $n \in \mathbb{N}$, be a sequence of states determined by a Markov chain.

For $t \in \mathbb{R}^+$, define the Markov chain perturbation $Z_t = X_{\lfloor t/h \rfloor}$, where $\lfloor t/h \rfloor$ is the greatest integer less than or equal to t/h .

Note that Z_t , instead of being defined only on discrete time values like a Markov chain, is instead a stepwise function defined on continuous time.

Definition 4. Given a metric space M and state space S as above, define a dynamical system φ with random perturbation function Z_t , as given in [Definition 3](#), by

$$\varphi : \mathbb{R}^+ \times M \times S \rightarrow M,$$

with

$$\varphi(t, x_0, Z_0) = \varphi_{Z_t}(t - nh, \varphi_{Z_{nh}}(h, \dots \varphi_{Z_{2h}}(h, \varphi_{Z_h}(h, \varphi_{Z_0}(h, x_0))))),$$

where φ_{Z_k} represents the deterministic dynamical system φ evaluated in state Z_k and nh is the largest multiple of h less than t .

For ease of notation, let

$$x_t = \varphi(t, x_0, Z_0) \in M$$

represent the position of the system at time t .

Definition 5. Let

$$Y_t = \begin{pmatrix} x_t \\ Z_t \end{pmatrix}$$

define the hybrid system at time t . In other words, the hybrid system consists of both a position $x_t = \varphi(t, x_0, Z_0) \in M$ and a state $Z_t \in S$.

The ω -limit set has the following generalization in a hybrid system.

Definition 6. The stochastic limit set $C(x)$ for an element of our state space $x \in M$ and the hybrid system given above is the subset of M with the following three properties:

- (1) Given $y \in M$ and $t_k \rightarrow \infty$ such that $x_{t_k} \rightarrow y$, $P(y \in C(x)) = 1$.
- (2) $C(x)$ is closed.
- (3) $C(x)$ is minimal: if some set $C'(x)$ has properties 1 and 2, then $C \subseteq C'$.

3. The hybrid system as a Markov process

Lemma 7. Each of the following is a Markov process:

- (i) Any deterministic dynamical system $\varphi(t, x_0)$.
- (ii) Any Markov chain perturbation Z_t , as in [Definition 2](#).
- (iii) The corresponding hybrid system Y_t , as in [Definition 5](#).

Proof. (i) Any deterministic system is trivially a Markov process, since $\varphi(t, x_0)$ is uniquely determined by $\varphi(\tau, x_0)$ at any single past time $\tau \in \mathbb{R}^+$.

(ii) By definition, a Markov chain is a Markov process. However, the Markov chain perturbation Z_t is not exactly a Markov chain. A Markov chain exists on a discrete time set, in our case given by $T = \{t \in \mathbb{R}^+ \mid t = nh \text{ for some } n \in \mathbb{N}\}$; conversely, the time set of Z_t is \mathbb{R}^+ , with transitions between states occurring on the previous time set (that is, at $t \equiv 0 \pmod{h}$). Despite this difference, Z_t maintains the Markov property: we can compute $P(Z_t \in A)$ for any set A based solely on Z_{τ_1} and the values of the times t and τ_1 . Explicitly, the probability that Z_t will be in state s_i at time t is given by

$$P(Z_t = s_i) = ((Q^T)^n)_{ij},$$

where n is the number of integer multiples of h (i.e., the number of transitions that occur) between t and τ_1 . Clearly, this is independent of the states Z_{τ_i} for $i > 1$, so that the random perturbation is indeed a Markov process.

(iii) Now, keeping in mind that the hybrid system Y_t consists of both a location $x_t \in M$ in the state space and a value $Z_t \in S$ of the random component, we can combine (i) and (ii) to see that the entire system is also a Markov process. We see from (ii) that Z_t follows a Markov process. Furthermore, $P(x_t \in A_x)$ at time t depends solely on the location x_{τ_1} at any time $\tau_1 < t$ and the states of the random perturbation sequence Z between t and τ_1 , regardless of any past behavior of the system. Hence, for any collection of sets A_α , $\alpha \in \mathbb{N}$,

$$P(Z_t \in A_z \mid Z_{\tau_1} \in A_{z_1}, Z_{\tau_2} \in A_{z_2}, \dots, Z_{\tau_n} \in A_{z_n}) = P(Z_t \in A_z \mid Z_{\tau_1} \in A_{z_1}),$$

$$P(x_t \in A_x \mid x_{\tau_1} \in A_{x_1}, x_{\tau_2} \in A_{x_2}, \dots, x_{\tau_n} \in A_{x_n}) = P(x_t \in A_x \mid x_{\tau_1} \in A_{x_1}).$$

So,

$$P(Y_t \in A_y \mid Y_{\tau_1} \in A_{y_1}, Y_{\tau_2} \in A_{y_2}, \dots, Y_{\tau_n} \in A_{y_n}) = P(Y_t \in A_y \mid Y_{\tau_1} \in A_{y_1}).$$

Thus, the hybrid system is a Markov process. \square

Unfortunately, the hybrid system is not time-homogeneous. Recall that state transitions of Z_t occur at times $t = nh$ for $n \in \mathbb{N}$. So, the state of the system at time $h/4$ uniquely determines the system at $3h/4$, since there is no transition in this interval. However, the system at time $5h/4$ is not determined uniquely by the system at $3h/4$, since a stochastic transition occurs at $t = h \in [\frac{3}{4}h, \frac{5}{4}h]$. Therefore, with $t_1 = h/4$, $t_2 = 3h/4$, and $k = \frac{1}{2}$,

$$P(Y_{\frac{h}{4} + \frac{1}{2}} \in A \mid Y_{\frac{h}{4}} \in A_0) \neq P(Y_{\frac{3h}{4} + \frac{1}{2}} \in A \mid Y_{\frac{3h}{4}} \in A_0),$$

violating [Definition 1](#). However, in order to satisfy the hypotheses of the Krylov–Bogolyubov theorem [[Hairer 2010; 2006](#)] found in [Theorem 14](#), the hybrid system must be time-homogeneous.

To create a time-homogeneous system, we restrict the time set on which our Markov process is defined. Instead of allowing our time set

$$\{t, \tau_1, \tau_2, \tau_3, \dots, \tau_n\} \subset \mathbb{R}^+$$

to be any decreasing sequence of real numbers, we create time sets $t_0 + nh$ for each $t_0 \in [0, h)$ and $n \in \mathbb{N}$. In other words, we define a different time set for each value $t_0 < h$ as

$$\{t \in \mathbb{R}^+ \mid t = t_0 + nh \text{ for some } n \in \mathbb{N}\}.$$

We call the hybrid system on these multiple, restricted time sets the *discrete system*.

Proposition 8. *The discrete hybrid system above is a time-homogeneous Markov process.*

Proof. First, we must show that the discrete hybrid system is a Markov process at all. This follows immediately from the proof that our original hybrid system is a Markov process. Since the Markov property holds for all $t, \tau_1, \tau_2, \dots, \tau_n \in \mathbb{R}^+$, it must necessarily hold for the specific time set

$$\{t \in \mathbb{R}^+ \mid \text{there exists } n \in \mathbb{N} \text{ such that } t = t_0 + nh\}$$

for each $t_0 < h$.

Now, it remains to show that this system is time-homogeneous. Recall that the time-continuous hybrid system failed to be time-homogeneous because its Z_t component was not time-homogeneous. Although transitions occurred only at regular, discrete time values, a test interval could be of any length; an interval of size $h/2$, for example, might contain either 0 or 1 transitions. However, because our discrete system creates separate time sets, any time interval—starting and ending within the same time set—must be of length nh for some $n \in \mathbb{N}$, and thus will contain precisely n potential transitions. So, taking $t_1, t_2 \in \mathbb{R}^+$, we know that

$$P(Y_{t_1+nh} \in A \mid Y_{t_1} \in A_0) = P(Y_{t_2+nh} \in A \mid Y_{t_2} \in A_0).$$

Note that the first component of the hybrid system, x_t , is also time-homogeneous under the discrete time system. Given Z_t , it can be treated as a deterministic system, and therefore time-homogeneous. Thus, the discrete hybrid system is time-homogeneous. \square

4. Invariant measures for the hybrid system

We now introduce several definitions that will lead to the main results of this paper.

Definition 9. Consider a hybrid system Y_t and a σ -algebra Σ on the space M . A measure μ on M is invariant if, for all sets $A \in \Sigma$ and all times $t \in \mathbb{R}^+$,

$$\mu(A) = \int_{x_0 \in M} P(x_t \in A) \mu(dx).$$

Definition 10. Let (M, \mathcal{T}) be a topological space, and let Σ be a σ -algebra on M that contains the topology \mathcal{T} . Let \mathcal{M} be a collection of probability measures defined on Σ . The collection \mathcal{M} is called tight if, for any $\epsilon > 0$, there is a compact subset K_ϵ of M such that, for any measure μ in \mathcal{M} ,

$$\mu(M \setminus K_\epsilon) < \epsilon.$$

Note that since μ is a probability measure, it is equivalent to, say, $\mu(K_\epsilon) > 1 - \epsilon$.

The following definitions are from [Hairer 2010].

Definition 11. Let (M, ρ) be a separable metric space. Let $\{\mathcal{P}(M)\}$ denote the collection of all probability measures defined on M (with its Borel σ -algebra).

A collection $K \subset \{\mathcal{P}(M)\}$ of probability measures is tight if and only if K is sequentially compact in the space equipped with the topology of weak convergence.

Definition 12. Consider M with σ -algebra Σ . Let $C^0(M, \mathbb{R})$ denote the set of continuous functions from M to \mathbb{R} . The probability measure $\mathcal{P}(t, x, \cdot)$ on Σ induces a map

$$\mathcal{P}_t(x) : C^0(M, \mathbb{R}) \rightarrow \mathbb{R}, \quad \text{with} \quad \mathcal{P}_t(x)(f) = \int_{y \in M} f(y) \mathcal{P}(t, x, dy).$$

\mathcal{P}_t is called a *Markov operator*.

Definition 13. A Markov operator \mathcal{P} is Feller if $\mathcal{P}\varphi$ is continuous for every continuous bounded function $\varphi : X \rightarrow \mathbb{R}$. In other words, it is Feller if and only if the map $x \mapsto \mathcal{P}(x, \cdot)$ is continuous in the topology of weak convergence.

We state the Krylov–Bogolyubov theorem without proof.

Theorem 14 (Krylov–Bogolyubov). *Let \mathcal{P} be a Feller Markov operator over a complete and separable space X . Assume that there exists $x_0 \in X$ such that the sequence $\mathcal{P}^n(x_0, \cdot)$ is tight. Then, there exists at least one invariant probability measure for \mathcal{P} .*

We now show that the conditions of the theorem are satisfied by the discrete hybrid system, yielding the existence of invariant measures as a corollary.

Lemma 15. *Given $t_0 \in [0, h)$, the discrete hybrid system Markov operators \mathcal{P}_n for $n \in \mathbb{N}$ given by*

$$\mathcal{P}_n f(Y) \equiv \int_{M \times S} f(Y_1) \mathcal{P}(nh, Y, dY_1)$$

are Feller.

Proof. We begin by showing that \mathcal{P}_1 is Feller. By induction, it follows that \mathcal{P}_n is Feller for all $n \in \mathbb{N}$. It is clear that there are only finitely many possible outcomes of running the hybrid system for time h . Namely, there are at most $|S|$ possible outcomes, where $|S|$ denotes the cardinality of S . Given

$$Y_0 = \begin{pmatrix} x_0 \\ Z_0 = s_i \end{pmatrix} \in M \times S,$$

the only possible outcomes at time $t = 1$ are

$$Y_1^j = \begin{pmatrix} \varphi_j(t_0, \varphi_i(h - t_0, x)) \\ s_j \end{pmatrix}$$

for $j \in \{1, \dots, |S|\}$, where φ_i, φ_j are the flows of the dynamical systems corresponding to states s_i and s_j , respectively. The probability of the j -th outcome is

given by $P_{i \rightarrow j}$, the probability of transitioning from state s_i to state s_j . Therefore,

$$\mathcal{P}_1 f(Y) = \int_{M \times S} f(Y_1) \mathcal{P}(h, Y, dY_1) = \sum_{j=1}^{|S|} P_{i \rightarrow j} f(Y_1^j).$$

Each φ_i is continuous under the assumption that each flow is continuous with respect to its initial conditions. The map from s_i to s_j is continuous since S is finite, so every set is open and hence the inverse image of any open set is open. The function f is continuous by hypothesis, and any finite sum of continuous functions is also continuous. Therefore $\mathcal{P}_1 f$ is also continuous, and hence \mathcal{P}_1 is Feller. \square

We see now that the conditions of [Theorem 14](#) (Krylov–Bogolyubov) hold. Namely, because M and S are compact (the former by assumption, the latter since it is finite), $M \times S$ is compact. Thus, any collection of measures is automatically tight, since we can take $K_\epsilon = X$. It is well known that any compact metric space is also complete and separable. Applying [Theorem 14](#), then, gives the following corollary, which is one of the primary results of the paper.

Corollary 16. *The discrete hybrid system has an invariant measure for each $t_0 \in [0, h)$.*

So, rather than speaking of an invariant measure for the time-continuous hybrid system, we can instead imagine a periodic invariant measure cycling continuously through h . That is, for each time $t_0 \in [0, h)$, there exists a measure μ_{t_0} such that for $t \equiv 0 \pmod{h}$,

$$\mu_{t_0}(A) = \int_{Y \in M \times S} \mathcal{P}(t, Y, A) d\mu_{t_0}.$$

The measure μ_{t_0} above is a measure on the product space $M \times S$, since this is where the hybrid system lives. However, what we are really after is an invariant measure on just M , the space where the dynamical system part of the hybrid system lives. Fortunately, we can define a measure on M by the following construction.

Proposition 17. *Given μ_t , an invariant probability measure on $M \times S$, the function*

$$\tilde{\mu}_t(A) \equiv \mu_t(A, S),$$

where $A \subseteq M$ is an invariant probability measure on M .

Proof. The fact that $\tilde{\mu}_t$ is a probability measure follows almost immediately from the fact that μ_t is a probability measure. The probability that $x_t \in \emptyset$ is 0, so $\tilde{\mu}_t(\emptyset) = 0$. The probability that $x_t \in M$ is 1, so $\tilde{\mu}_t(M) = 1$. Countable additivity of $\tilde{\mu}_t$ follows from countable additivity of μ_t . Therefore, $\tilde{\mu}_t$ is a probability measure on M . \square

Thus far, we have proven the existence of a measure μ_{t_0} for $t_0 \in [0, h)$ such that for $t \equiv 0 \pmod{h}$,

$$\mu_{t_0}(A) = \int_{x_0 \in M, s \in S} P(\varphi(t, x_0, s) \in A) d\mu_{t_0}.$$

The following theorem relates the collection of invariant measures $\{\tilde{\mu}_{t_0}\}$ using the flow φ . This is the main result of the paper.

Theorem 18. *Given invariant measure μ_0 , the measure μ_t defined by*

$$\mu_t(A) = \sum_{s \in S} \int_{x_0 \in M} P(\varphi(t, x_0, s) \in A) d\mu_0$$

is also invariant in the sense that $\mu_t = \mu_{t+nh}$ for $n \in \mathbb{N}$.

Proof. We will show that $\mu_t = \mu_{t+h}$. By induction, this implies that $\mu_t = \mu_{t+nh}$ for all $n \in \mathbb{N}$. We have

$$\mu_{t+h}(A) = \sum_{s \in S} \int_{x_0 \in M} P(\varphi(t+h, x_0, s) \in A) d\mu_0.$$

Applying the definition of conditional probability,

$$\begin{aligned} & \sum_{s \in S} \int_{x_0 \in M} P(\varphi(t+h, x_0, s) \in A) d\mu_0 \\ &= \sum_{r \in S} \int_{y \in M} \left[P(\varphi(t, y, r) \in A) \sum_{s \in S} \int_{x_0 \in M} P(\varphi(h, x_0, s) \in dy \times \{r\}) d\mu_0 \right]. \end{aligned}$$

Loosely speaking, the probability that a trajectory beginning at (x, s) will end in a set A after a time $t+h$ is the product of the probability that a trajectory beginning at (y, r) will end in A after a time t multiplied by the probability that a trajectory beginning at (x, s) will end at (y, r) after a time h , integrating over all possible pairs (y, r) . Here, we have implicitly used the fact that the hybrid system is a Markov process to ensure that the state of the system at time $t+h$ given the state at time h is independent of the initial state, and we have avoided the problem of time-inhomogeneity by considering trajectories that only begin at times congruent to $0 \pmod{h}$.

Furthermore, we have

$$\mu_h(dy \times \{r\}) = \sum_{s \in S} \int_{x_0 \in M} P(\varphi(h, x_0, s) \in dy \times \{r\}) d\mu_0$$

and

$$\mu_h(dy \times \{r\}) = d\mu_h(y, r);$$

so,

$$\mu_{t+h}(A) = \sum_{r \in S} \int_{y \in M} P(\varphi(t, y, r) \in A) d\mu_h.$$

Since μ_0 is invariant by assumption, $\mu_0 = \mu_h$. Therefore,

$$\mu_{t+h}(A) = \sum_{r \in S} \int_{y \in M} P(\varphi(t, y, r) \in A) d\mu_0 = \mu_t(A). \quad \square$$

5. Examples

Some examples of hybrid systems can be found in [Ayers 2010; Baldwin 2007]. Here, we will examine two simple cases to illustrate the theory developed above.

5.1. A one-dimensional hybrid system. We begin with a one-dimensional linear dynamical system with a stochastic perturbation:

$$\dot{x} = -x + Z_t,$$

where $Z_t \in \{-1, 1\}$. Both components of this system have a single, attractive equilibrium point: for $Z_t = 1$, this is $x = 1$, and for $Z_t = -1$, $x = -1$. At timesteps of length $h = 1$, Z_t is perturbed by a Markov chain given by the transition matrix Q . Q is therefore a 2×2 matrix of nonnegative entries,

$$Q = \begin{pmatrix} P_{1 \rightarrow 1} & P_{1 \rightarrow -1} \\ P_{-1 \rightarrow 1} & P_{-1 \rightarrow -1} \end{pmatrix},$$

where $P_{i \rightarrow j}$ gives the probability of the equilibrium point transitioning from i to j at each integer timestep. Since the total probability measure must equal 1,

$$\sum_{j \in \{1, -1\}} P_{i \rightarrow j} = 1, \quad i \in \{1, -1\}.$$

Furthermore, to avoid the deterministic case, we take $P_{i \rightarrow j} \neq 0$ for all i, j .

Proposition 19. *The stochastic limit set $C(x_0) = [-1, 1]$ for all $x_0 \in \mathbb{R}$.*

Proof. We begin by showing that $C(x) \subset [-1, 1]$: that is, that every possible trajectory in our system will eventually enter and never leave $[-1, 1]$, meaning that no it is only possible to have $t^* \rightarrow \infty$ such that $x^* = y$ for $y \in [-1, 1]$. First, consider $x_0 \in [-1, 1]$. If we are in state $Z_t = 1$, then the trajectory is attracted upwards and bounded above by $x = 1$; in state $Z_t = -1$, the trajectory is attracted downwards and bounded below by $x = -1$. In both cases, the trajectory cannot move above 1 or below -1 , and so will remain in $[-1, 1]$ for all time.

Now, consider $x_0 \notin [-1, 1]$. If the trajectory ever enters $[-1, 1]$, by similar argument as above, it will remain in that region for all time. So, it remains to show that $\varphi(t, x_0, Z_0) \in [-1, 1]$ for some $t \in \mathbb{R}$. First, take $x_0 > 1$. In either state,

the trajectory will be attracted downward, and will eventually enter $[1, 2]$ at time t_2 . Once there, at the first timestep in which $Z_t = -1$ it will cross $x = 1$ and enter $[-1, 1]$. And since we have taken all entries of the transition probability matrix Q to be nonzero, there almost surely exists a time $t_3 > t_2$ for which the state is $Z_t = -1$; then, the trajectory will enter $[-1, 1]$ and never leave. By similar argument, any trajectory starting at $x_0 < -1$ will enter and never leave $[-1, 1]$. Thus, $C(x) \subset [-1, 1]$.

Now, we must show that $[-1, 1] \in C(x)$: that is, that for every trajectory $\varphi(t, x_0, Z_0)$ and every point $y \in [-1, 1]$ there is $t^* \rightarrow \infty$ such that $\varphi(t^*, x_0, Z_0) \rightarrow y$. To do this, we really only need to show that given any point $x_0 \in [-1, 1]$ and any transition matrix Q , there almost surely exists some time t^* with $\varphi(t^*, x_0, Z_0) = x^*$. If one such time t^* is guaranteed to exist, then we can iterate the process for a solution beginning at (t^*, x^*) to produce an infinite sequence of times. To show that t^* exists, we calculate a lower bound on the probability that $\varphi(t_n, x_0, Z_0) = x^*$.

Without loss of generality, suppose that $x_0 > x^*$. We have already shown that any solution will enter $[-1, 1]$, so take $\sup(x_0) = 1$. From here, we can calculate the minimum number of necessary consecutive periods, k , for which $Z_n = -1$ in order for a solution with $x_0 = 1$ to decay to x^* . The probability of this sequence of k consecutive periods occurring is given by

$$P_1(k) = (P_{1 \rightarrow -1})(P_{-1 \rightarrow -1})^{k-1}$$

if $Z_0 = 1$ and

$$P_{-1}(k) = (P_{-1 \rightarrow -1})^k$$

if $Z_0 = -1$. Thus, for some $t^* \in [0, k]$,

$$P(\varphi(t^*, x_0, Z_0) = x^*) \geq \min(P_1(k), P_{-1}(k)) > 0,$$

since $P_{i \rightarrow j} > 0$. So,

$$P(t^* \notin [0, k]) \leq 1 - P(x^*) < 1 \quad \text{and} \quad P(t^* \notin [0, mk]) \leq (1 - P(x^*))^m.$$

As $m \rightarrow \infty$, $(1 - P(x^*))^m \rightarrow 0$. So, with probability 1, there exists t^* with $\varphi(t^*, x_0, Z_0) = x^*$.

By similar argument, for $x_0 < x^*$ and all $x^* \in (-1, 1)$, we can find a time sequence $\{t_n\}$ such that $\varphi(t_n, x_0, Z_0) = x^*$. So, we know that for all $x^* \in (-1, 1)$, $x^* \in C(x)$.

So, we have proven that $[-1, 1] \subseteq C(x)$ and $(-1, 1) \subseteq C(x)$. Since $C(x)$ must by definition be closed, $C(x) = [-1, 1]$. \square

We can study the behavior of this system numerically. **Figure 1** (left) depicts a solution calculated for the transition matrix

$$Q_1 = \begin{pmatrix} 0.4 & 0.6 \\ 0.5 & 0.5 \end{pmatrix},$$

with initial values $x_0 = 2, Z_0 = 1$.

As expected, the trajectory enters the interval $(-1, 1)$ and stays there for all time, oscillating between $x = -1$ and $x = 1$. Intuitively, it seems that the trajectory will cross any x^* in this interval repeatedly, so that indeed $C(x) = [-1, 1]$. This is not quite so clear for the transition matrix

$$Q_2 = \begin{pmatrix} 0.1 & 0.9 \\ 0.1 & 0.9 \end{pmatrix},$$

which yields the trajectory shown in **Figure 1** (right) for $x_0 = 2, Z_0 = 1$.

It may appear that some set of points near $x = 1$ might be crossed by our path only a finite number of times. But, as proven above, any point in $(-1, 1)$ will almost surely be reached infinitely many times as $t \rightarrow \infty$, so $C(x) = [-1, 1]$.

Now, we consider the eigenvalues and eigenvectors of the transition matrices. The eigenvector of Q_1^T with eigenvalue 1 is

$$\vec{v} = \begin{pmatrix} \frac{5}{11} \\ \frac{6}{11} \end{pmatrix},$$

and the eigenvector of Q_2^T with eigenvalue 1 is

$$\vec{v}' = \begin{pmatrix} \frac{9}{10} \\ \frac{1}{10} \end{pmatrix}.$$

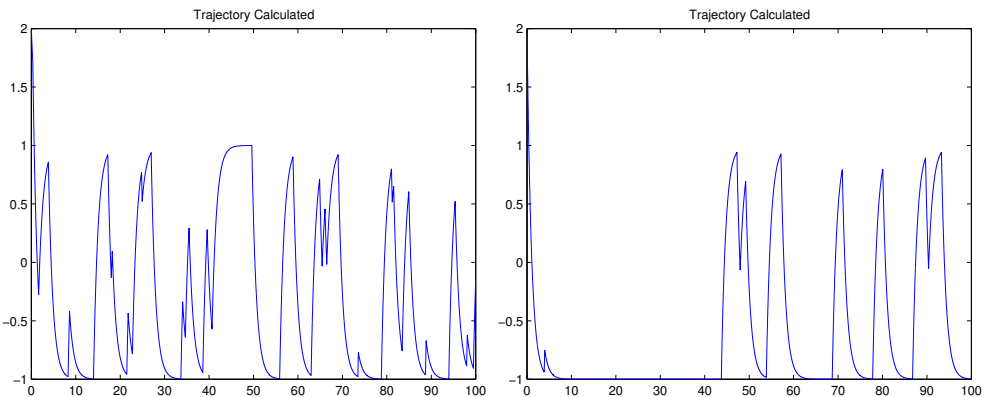


Figure 1. A sample trajectory for a hybrid system with transition matrix Q_1 (left) and Q_2 (right).

These eigenvectors give the invariant measures on the state space S . We know from [Proposition 17](#) that there also exists an invariant measure on M . Here, since any trajectory in M will almost surely enter $C(x) = [-1, 1]$, the support of the invariant measure must be contained in $C(x)$. It is not difficult to see that this invariant measure cannot be constant for all $t \in \mathbb{R}^+$. Given any point $x_0 \in [-1, 1]$, we know that at $t = 1$, one of two things will have happened to the trajectory:

- (i) it will have decayed exponentially toward $x = 1$, if $Z_1 = 1$, or
- (ii) it will have decayed exponentially toward $x = -1$, if $Z_1 = -1$.

In case (i), if a solution begins at $x_0 = -1$ for $t = 0$, the solution will have decayed to a value of $1 - 2e^{-1} \approx 0.264$ by $t = 1$. In case (ii) a solution beginning at $x_0 = 1$ for $t = 0$ will decay to a value of $-1 + 2e^{-1} \approx -0.264$. Thus, if we are in case (i), all trajectories in $[-1, 1]$ at $t = n$ will be located in $[0.264, 1]$ at $t = n + 1$. If we are in case (ii), all will be in $[-1, -0.264]$. It is not possible for any trajectory to be located in $[-0.264, 0.264]$ at an integer time value. But, clearly, some solutions will cross into this region, as depicted in [Figure 2](#). Therefore, no probability distribution will remain constant for all t in the time set \mathbb{R}^+ .

However, as [Figure 2](#) suggests, there is some distribution that is invariant under $t \rightarrow t + n$ for $n \in \mathbb{N}$. Approximations of the invariant measures at $t \in [0, 1]$ for transition matrix Q_1 are shown in [Figure 3](#).

5.2. A two-dimensional hybrid system. Our second example is a two-dimensional system used to model the kinetics of chemical reactors. The general system $f(x_1, x_2)$ is given by

$$\begin{aligned}\dot{x}_1 &= -\lambda x_1 - \beta(x_1 - x_c) + B Da f(x_1, x_2), \\ \dot{x}_2 &= -\lambda x_2 + Da f(x_1, x_2),\end{aligned}$$

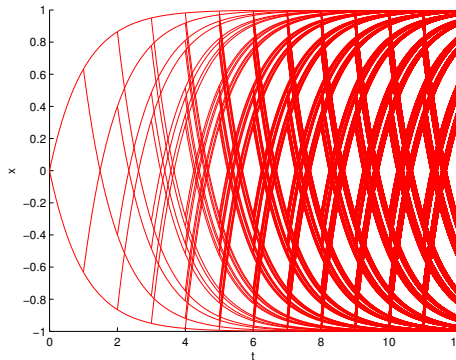


Figure 2. A spider plot showing all possible trajectories starting at $x_0 = 0$.

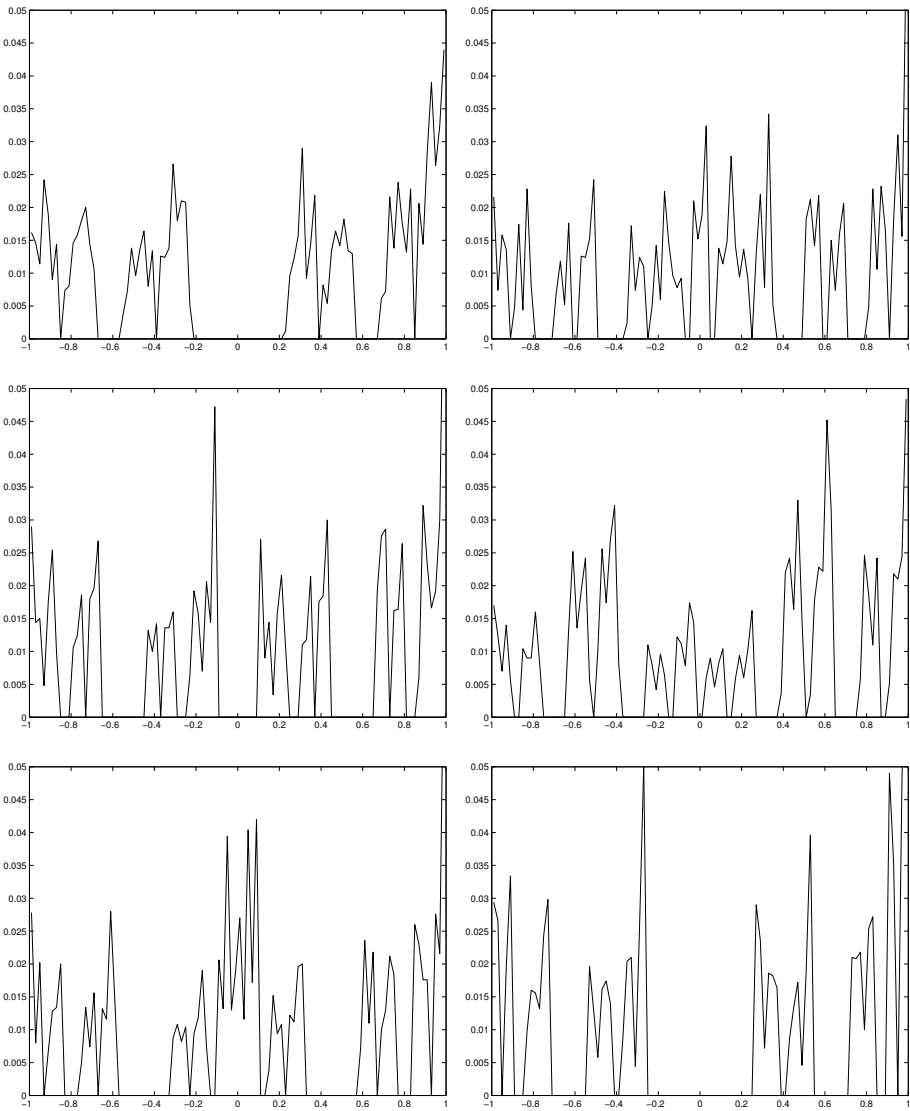


Figure 3. The invariant measure $\tilde{\mu}_0$ for a hybrid system with transition matrix Q_1 .

where λ , β , x_c , Da , and B are physical parameters (see [Poore 1973]). Here, we use a simplified application of the system:

$$\begin{aligned} \dot{x}_1 &= -x_1 - 0.15(x_1 - 1) + 0.35(1 - x_2)e^{x_1} + Z_t(1 - x_1), \\ \dot{x}_2 &= -x_2 + 0.05(1 - x_2)e^{x_1}. \end{aligned}$$

This system is used to describe a continuous stirred tank reactor (CSTR). This type

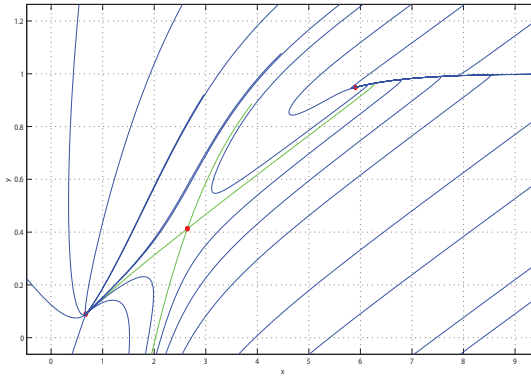


Figure 4. Phase plane of the deterministic system, $Z_n = 0$.

of reactor is used to control chemical reactions that require a continuous flow of reactants and products and are easy to control the temperature with. They are also useful for reactions that require working with two phases of chemicals.

To understand the behavior of this system mathematically, we set our stochastic variable $Z_t = 0$ and treat it as a deterministic system. This system has three fixed points, approximately at $(0.67, 0.09)$, $(2.64, 0.41)$, and $(5.90, 0.95)$; the former and latter are attractor points, while the middle is a saddle point, as shown in Figure 4. The saddle point $(2.64, 0.41)$ creates a separatrix, a repelling equilibrium line between the two attracting fixed points. These points, $(0.67, 0.09)$ and $(5.90, 0.95)$, comprise the ω -limit set of our state space.

With this information, we proceed to analyze the stochastic system. As discussed above, the random variable here is Z_t , which in applications can take values between -0.15 and 0.15 . To understand the full variability of this system, we take

$$Z_t \in \{-0.15, 0, 0.15\}$$

with the transition matrix

$$\begin{pmatrix} 0.3 & 0.3 & 0.4 \\ 0.3 & 0.3 & 0.4 \\ 0.3 & 0.3 & 0.4 \end{pmatrix},$$

yielding the phase plane in Figure 5.

We see that, for x_0 away from the separatrices, $\varphi(t, x_0, Z_0)$ behaves similarly to $\varphi(t, x_0)$. Although state changes create some variability in a given trajectory, these paths move toward the groups of associated attracting fixed points, which define the stochastic limit sets for this system. However, $\varphi(t, x_0, Z_0)$ for x_0 between the red and green separatrices is unpredictable; depending on the sequence of state changes for a given trajectory, it might move either to the right or the left of the region

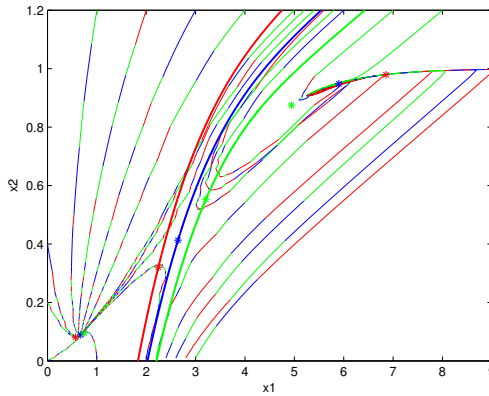


Figure 5. Phase plane with randomness, showing fixed points, separatrices, and portions of trajectories. Red, blue and green indicate states 1 ($Z_t = -0.15$), 2 ($Z_t = 0$) and 3 ($Z_t = 0.15$).

defined by the separatrices. This area is the *bistable region*, because a trajectory beginning within it has two separate stochastic limit sets.

For example, we have in [Figure 6](#) a spider plot beginning in the bistable region at (3.5, 0.75). A spider plot shows all possible trajectories starting from a single point in a hybrid system by, at each timestep, taking every possible state.

Thus, we see that the introduction of a stochastic element to a deterministic system can grossly affect the outcome of the system, as a trajectory can now cross any of the separatrices by being in a different state.

The stochastic element also affects the behavior of the hybrid system around the invariant region. In [Figure 7](#), we show the path of a single trajectory in the invariant region defined by the fixed points near (0.67, 0.9). Plotting this trajectory for a long period of time approximates the invariant region that would appear if we ran a spider plot from the same point, but much more clearly.

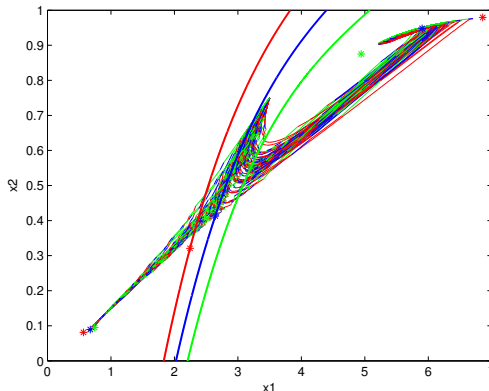


Figure 6. Spider plot. Color scheme as in [Figure 5](#).

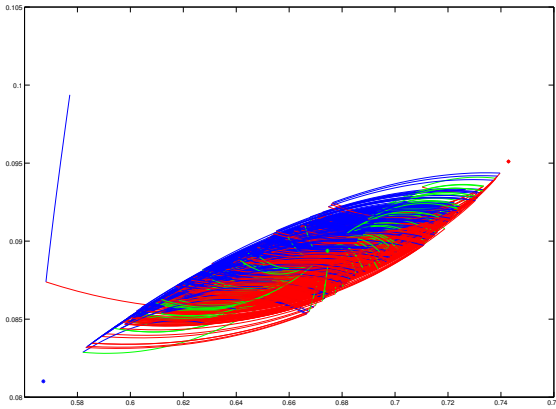


Figure 7. Random trajectory.

As we saw in the one-dimensional system, considering the counts taken at specific times in the interval between two state changes, $h = 1$ (since our state transitions occur on \mathbb{N}), yields a periodic set of invariant measures. Similarly to [Figure 3](#), [Figure 8](#) shows the positions of our random trajectory in the invariant region at time $t, \text{ mod } h$.

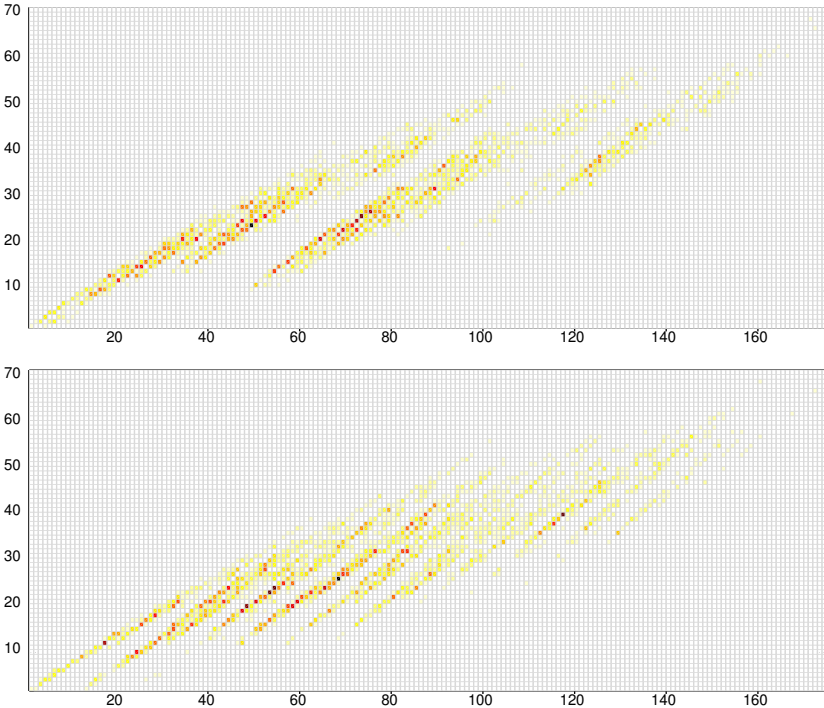


Figure 8. Count of trajectory paths within one timestep.

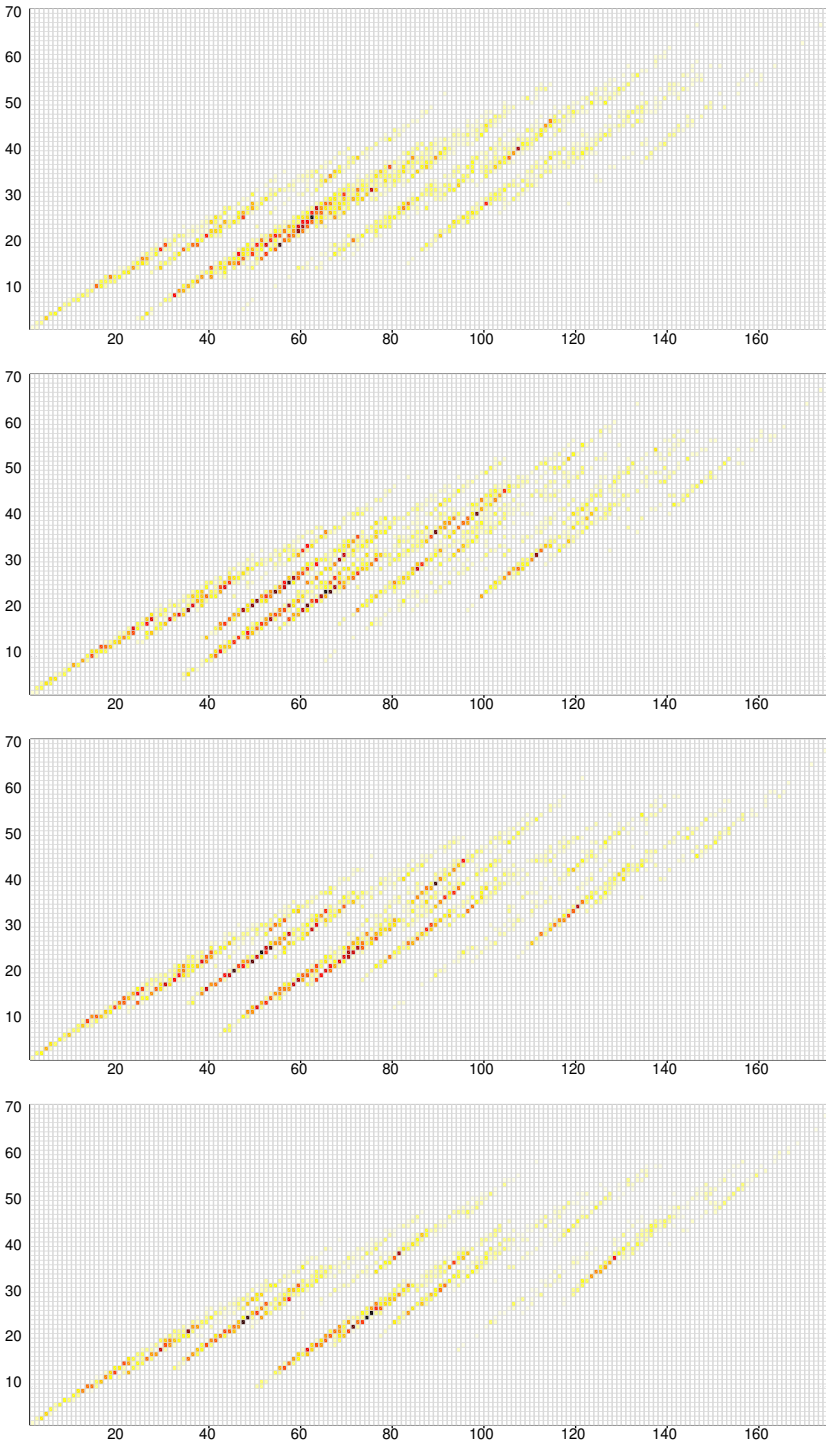


Figure 8. Count of trajectory paths within one timestep (continued).

A denser series of count images would show more clearly that the invariant measure at $t \bmod h$ cycles continuously.

6. Conclusion

We have studied hybrid systems consisting of a finite set S of dynamical systems over a compact space M with a Markov chain on S acting at discrete time intervals. Such a hybrid system is a Markov process, which can be made time-homogeneous by discretizing the system. Then, there exists a family of invariant measures on the product space $M \times S$, which can be projected onto a family of measures on M . We have demonstrated a relation between the members of this family.

We have studied both a one-dimensional and a two-dimensional example of a hybrid system. These examples provide insight into the stochastic equivalent of ω -limit sets and yield graphical representations of the invariant measures on these sets.

Acknowledgements

We wish to recognize Kimberly Ayers for her helpful discussions and Professor Wolfgang Kliemann for his instruction and guidance. We would like to thank the Department of Mathematics at Iowa State University for their hospitality during the completion of this work. In addition, we would like to thank Iowa State University, Alliance, and the National Science Foundation for their support of this research. Figure 4 was drawn using the “pplane8.m” Matlab program.

References

- [Ayers 2010] K. D. Ayers, “Stochastic perturbations of the Fitzhugh–Nagumo equations”, Undergraduate honors thesis, Bowdoin College, 2010.
- [Baldwin 2007] M. C. Baldwin, “Stochastic analysis of Marotzke and Stone climate model”, Master’s thesis, Iowa State University, 2007.
- [Hairer 2006] M. Hairer, “Ergodic properties of Markov processes”, lecture notes, 2006, available at <http://www.hairer.org/notes/Markov.pdf>.
- [Hairer 2010] M. Hairer, “Convergence of Markov processes”, lecture notes, 2010, available at <http://www.hairer.org/notes/Convergence.pdf>.
- [Poore 1973] A. B. Poore, “A model equation arising from chemical reactor theory”, *Arch. Rational Mech. Anal.* **52** (1973), 358–388. MR 49 #3272

Received: 2013-07-16

Accepted: 2013-10-05

garci363@umn.edu

*Department of Mathematics, University of Minnesota,
Minneapolis, MN 55455, United States*

jckunze@smcm.edu

*Mathematics and Computer Science Department, St. Mary’s
College of Maryland, St. Mary’s City, MD 20686, United States*

- twr27@cornell.edu *Department of Mathematics, Cornell University,
Ithaca, NY 14850, United States*
- anthony.sanchez.1@asu.edu *School of Mathematical and Statistical Sciences, Arizona
State University, Tempe, AZ 85287-1804, United States*
- sshao@iastate.edu *Department of Mathematics, Iowa State University,
Ames, IA 50011, United States*
- esperanza@carroll.edu *Department of Mathematics, Engineering, and Computer
Science, Carroll College, Helena, MT 59625, United States*
- cvidden@iastate.edu *Department of Mathematics, Iowa State University,
Ames, IA 50011, United States*

EDITORS

MANAGING EDITOR

Kenneth S. Berenhaut, Wake Forest University, USA, berenhks@wfu.edu

BOARD OF EDITORS

Colin Adams	Williams College, USA colin.c.adams@williams.edu	David Larson	Texas A&M University, USA larson@math.tamu.edu
John V. Baxley	Wake Forest University, NC, USA baxley@wfu.edu	Suzanne Lenhart	University of Tennessee, USA lenhart@math.utk.edu
Arthur T. Benjamin	Harvey Mudd College, USA benjamin@hmc.edu	Chi-Kwong Li	College of William and Mary, USA ckli@math.wm.edu
Martin Bohner	Missouri U of Science and Technology, USA bohner@mst.edu	Robert B. Lund	Clemson University, USA lund@clemson.edu
Nigel Boston	University of Wisconsin, USA boston@math.wisc.edu	Gaven J. Martin	Massey University, New Zealand g.j.martin@massey.ac.nz
Amarjit S. Budhiraja	U of North Carolina, Chapel Hill, USA budhiraj@email.unc.edu	Mary Meyer	Colorado State University, USA meyer@stat.colostate.edu
Pietro Cerone	La Trobe University, Australia P.Cerone@latrobe.edu.au	Emil Minchev	Ruse, Bulgaria eminchev@hotmail.com
Scott Chapman	Sam Houston State University, USA scott.chapman@shsu.edu	Frank Morgan	Williams College, USA frank.morgan@williams.edu
Joshua N. Cooper	University of South Carolina, USA cooper@math.sc.edu	Mohammad Sal Moslehian	Ferdowsi University of Mashhad, Iran moslehian@ferdowsi.um.ac.ir
Jem N. Corcoran	University of Colorado, USA corcoran@colorado.edu	Zuhair Nashed	University of Central Florida, USA znashed@mail.ucf.edu
Toka Diagana	Howard University, USA tdiagana@howard.edu	Ken Ono	Emory University, USA ono@mathcs.emory.edu
Michael Dorff	Brigham Young University, USA mdorff@math.byu.edu	Timothy E. O'Brien	Loyola University Chicago, USA tbriell@luc.edu
Sever S. Dragomir	Victoria University, Australia sever@matilda.vu.edu.au	Joseph O'Rourke	Smith College, USA orourke@cs.smith.edu
Behrouz Emamizadeh	The Petroleum Institute, UAE bemamizadeh@pi.ac.ae	Yuval Peres	Microsoft Research, USA peres@microsoft.com
Joel Foisy	SUNY Potsdam foisyjs@potsdam.edu	Y.-F. S. Pétermann	Université de Genève, Switzerland petermann@math.unige.ch
Errin W. Fulp	Wake Forest University, USA fulp@wfu.edu	Robert J. Plemmons	Wake Forest University, USA rplemmons@wfu.edu
Joseph Gallian	University of Minnesota Duluth, USA kgallian@d.umn.edu	Carl B. Pomerance	Dartmouth College, USA carl.pomerance@dartmouth.edu
Stephan R. Garcia	Pomona College, USA stephan.garcia@pomona.edu	Vadim Ponomarenko	San Diego State University, USA vadim@sciences.sdsu.edu
Anant Godbole	East Tennessee State University, USA godbole@etsu.edu	Bjorn Poonen	UC Berkeley, USA poonen@math.berkeley.edu
Ron Gould	Emory University, USA rg@mathcs.emory.edu	James Propp	U Mass Lowell, USA jpropp@cs.uml.edu
Andrew Granville	Université Montréal, Canada andrew@dms.umontreal.ca	József H. Przytycki	George Washington University, USA przytyck@gwu.edu
Jerrold Griggs	University of South Carolina, USA griggs@math.sc.edu	Richard Rebarber	University of Nebraska, USA rrebarbe@math.unl.edu
Sat Gupta	U of North Carolina, Greensboro, USA sgupta@uncg.edu	Robert W. Robinson	University of Georgia, USA rwr@cs.uga.edu
Jim Haglund	University of Pennsylvania, USA jhaglund@math.upenn.edu	Filip Saidak	U of North Carolina, Greensboro, USA f_saidak@uncg.edu
Johnny Henderson	Baylor University, USA johnny_henderson@baylor.edu	James A. Sellers	Penn State University, USA sellersj@math.psu.edu
Jim Hoste	Pitzer College jhoste@pitzer.edu	Andrew J. Sterge	Honorary Editor andy@ajsterge.com
Natalia Hritonenko	Prairie View A&M University, USA nhritonenko@pvamu.edu	Ann Trenk	Wellesley College, USA atrenk@wellesley.edu
Glenn H. Hurlbert	Arizona State University, USA hurlbert@asu.edu	Ravi Vakil	Stanford University, USA vakil@math.stanford.edu
Charles R. Johnson	College of William and Mary, USA crjohnso@math.wm.edu	Antonia Vecchio	Consiglio Nazionale delle Ricerche, Italy antonia.vecchio@cnr.it
K. B. Kulasekera	Clemson University, USA kk@ces.clemson.edu	Ram U. Verma	University of Toledo, USA verma99@msn.com
Gerry Ladas	University of Rhode Island, USA gladas@math.uri.edu	John C. Wierman	Johns Hopkins University, USA wierman@jhu.edu
		Michael E. Zieve	University of Michigan, USA zieve@umich.edu

PRODUCTION


Silvio Levy, Scientific Editor

See inside back cover or msp.org/involve for submission instructions. The subscription price for 2014 is US \$120/year for the electronic version, and \$165/year (+\$35, if shipping outside the US) for print and electronic. Subscriptions, requests for back issues from the last three years and changes of subscribers address should be sent to MSP.

Involve (ISSN 1944-4184 electronic, 1944-4176 printed) at Mathematical Sciences Publishers, 798 Evans Hall #3840, c/o University of California, Berkeley, CA 94720-3840, is published continuously online. Periodical rate postage paid at Berkeley, CA 94704, and additional mailing offices.

Involve peer review and production are managed by EditFLOW[®] from Mathematical Sciences Publishers.

PUBLISHED BY

 **mathematical sciences publishers**
nonprofit scientific publishing

<http://msp.org/>

© 2014 Mathematical Sciences Publishers

involve

2014

vol. 7

no. 4

- [Whitehead graphs and separability in rank two](#) 431
MATT CLAY, JOHN CONANT AND NIVETHA RAMASUBRAMANIAN
- [Perimeter-minimizing pentagonal tilings](#) 453
PING NGAI CHUNG, MIGUEL A. FERNANDEZ, NIRALEE SHAH, LUIS SORDO VIEIRA AND ELENA WIKNER
- [Discrete time optimal control applied to pest control problems](#) 479
WANDI DING, RAYMOND HENDON, BRANDON CATHEY, EVAN LANCASTER AND ROBERT GERMICK
- [Distribution of genome rearrangement distance under double cut and join](#) 491
JACKIE CHRISTY, JOSH MCHUGH, MANDA RIEHL AND NOAH WILLIAMS
- [Mathematical modeling of integrin dynamics in initial formation of focal adhesions](#) 509
AURORA BLUCHER, MICHELLE SALAS, NICHOLAS WILLIAMS AND HANNAH L. CALLENDER
- [Investigating root multiplicities in the indefinite Kac–Moody algebra \$E_{10}\$](#) 529
VICKY KLIMA, TIMOTHY SHATLEY, KYLE THOMAS AND ANDREW WILSON
- [On a state model for the \$SO\(2n\)\$ Kauffman polynomial](#) 547
CARMEN CAPRAU, DAVID HEYWOOD AND DIONNE IBARRA
- [Invariant measures for hybrid stochastic systems](#) 565
XAVIER GARCIA, JENNIFER KUNZE, THOMAS RUDELIUS, ANTHONY SANCHEZ, SIJING SHAO, EMILY SPERANZA AND CHAD VIDDEN