A generalization of Zeckendorf's theorem
via circumscribed m-gons

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(Communicated by Stephan Garcia)

Zeckendorf's theorem states that every positive integer can be uniquely decomposed as a sum of nonconsecutive Fibonacci numbers, where the Fibonacci numbers satisfy $F_1 = 1$, $F_2 = 2$, and $F_n = F_{n-1} + F_{n-2}$ for $n \geq 3$. The distribution of the number of summands in such a decomposition converges to a Gaussian, the gaps between summands converge to geometric decay, and the distribution of the longest gap is similar to that of the longest run of heads in a biased coin; these results also hold more generally, though for technical reasons previous work is needed to assume the coefficients in the recurrence relation are nonnegative and the first term is positive.

We extend these results by creating an infinite family of integer sequences called the m-gonal sequences arising from a geometric construction using circumscribed m-gons. They satisfy a recurrence where the first $m+1$ leading terms vanish, and thus cannot be handled by existing techniques. We provide a notion of a legal decomposition, and prove that the decompositions exist and are unique. We then examine the distribution of the number of summands used in the decompositions and prove that it displays Gaussian behavior. There is geometric decay in the distribution of gaps, both for gaps taken from all integers in an interval and almost surely in distribution for the individual gap measures associated to each integer in the interval. We end by proving that the distribution of the longest gap between summands is strongly concentrated about its mean, behaving similarly as in the longest run of heads in tosses of a coin.

1. Introduction

The Fibonacci numbers are a heavily studied sequence which arises in many different ways and places. By defining them as $F_1 = 1$, $F_2 = 2$ and $F_{n+1} = F_n + F_{n-1}$, we have the remarkable property that every positive integer can be uniquely written as a sum of nonconsecutive Fibonacci numbers; further, this property is equivalent

to the definition of the Fibonacci numbers (i.e., if $\{a_n\}$ is a sequence of numbers such that every integer can be written uniquely as a sum of nonadjacent terms in the sequence, then $\{a_n\} = \{F_n\}$). Zeckendorf [1972] proved this in 1939, though he did not publish this result until much later.

In recent years many have studied generalizations to Zeckendorf's theorem by exploring different notions of decompositions and the properties of the associated sequences; see among others [Alpert 2009; Daykin 1960; Demontigny et al. 2014a; 2014b; Drmota and Gajdosik 1998; Filipponi et al. 1994; Grabner and Tichy 1990; Grabner et al. 1994; Keller 1972; Lengyel 2006; Miller and Wang 2012; 2014; Steiner 2002; 2005]. Despite the vast literature in this area, the majority of the research on generalized Zeckendorf decompositions has involved sequences with *positive linear recurrences*. Positive linear recurrence sequences $\{G_n\}$ satisfy a linear recurrence relation where the coefficients are nonnegative with the first and last term coefficients being positive.[1]

There has been little research which considers cases where the leading coefficient in the recurrence is zero; one such case is found in [Catral et al. 2014]. They studied what they call the Kentucky sequence, which is defined by the recurrence relation $H_{n+1} = H_{n-1} + 2H_{n-3}$ and $H_i = i$ for $i \leq 4$. While the behavior there is similar to the positive linear recurrences, there are sequences with very different behavior. One such is the Fibonacci quilt, which arises from creating a decomposition rule from the Fibonacci spiral[2] (see [Catral et al. 2016a; 2016b]), where the number of decompositions is not unique but in fact grows exponentially. This leads to the major motivation of this paper (as well as the motivation for the three papers just mentioned): how important is the assumption that the leading term is positive? The work mentioned above shows that it is not just a technically convenient assumption; markedly different behavior can emerge. Our goal is to try and determine when we have each type of behavior, and thus the purpose of this paper is to explore infinitely many recurrences with absent leading term and see the effect that has on the properties of the decompositions.

Specifically, we consider an infinite family of integer sequences called the *m-gonal sequences*, where $m \geq 3$. These sequences arise from a geometric construction using circumscribed m-gons, and after defining them below we state our results.

Definition of the m-gonal sequence.

One interpretation of Zeckendorf's theorem, which states that every positive integer can be written uniquely as a sum of nonconsecutive Fibonacci numbers, is that we have infinitely many bins with just one

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[1] Thus $G_{n+1} = c_1 G_n + \cdots + c_L G_{n-(L-1)}$ with $c_1 c_L > 0$ and $c_i \geq 0$.

[2] Let $f_n = F_{n-1}$, the standard definition of the Fibonacci numbers. Then the plane can be tiled in a spiral where the dimensions of the n-th square are $f_n \times f_n$; we declare a decomposition legal if no two summands used share an edge.
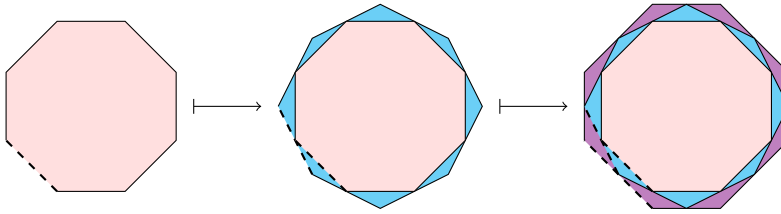
Figure 1. Circumscribed m-gons.

number per bin, and if we choose a bin to contribute a summand to a number's decomposition then we cannot choose a summand from an adjacent bin. We can generalize to bins with more elements, as well as disallow two bins to be used if they are within a given distance (see [Catral et al. 2014; 2016a; 2016b]). The m-gonal sequences are similar to these constructions, but have a two-dimensional structure arising from circumscribing m-gons about one central m-gon.

Briefly we view the decomposition rule corresponding to the m-gonal sequence for $m \geq 1$ by saying the sequence is partitioned into bins b_i of length $|b_i|$, where $|b_0| = 1$ and $|b_i| = m$ for all $i \geq 1$. A valid decomposition has no two summands being elements from the same bin. We refer to this decomposition as a *legal m-gonal decomposition of a positive integer z*. We now give details and examples of this construction.

The m-gonal sequence was initially constructed by circumscribing m-gons. For $m \geq 3$ we let M_0 denote a regular m-gon. Circumscribe the m-gon M_1 onto M_0 such that the vertices of M_0 bisect the edges of M_1. Note that this adds m faces to the resulting figure. We continue this process indefinitely, where we circumscribe the m-gon M_i onto M_{i-1} such that the vertices of M_{i-1} bisect the edges of M_i. At each step we have added an additional m faces to the resulting figure. We depict these initial iterations in Figure 1. Let $\mathcal{M}$ denote all of the faces created through the process of circumscribing m-gons. Then

$$\mathcal{M} = \{f_0\} \cup \left(\bigcup_{i=1}^{\infty} \{f(i, 1), f(i, 2), \ldots, f(i, m)\} \right), \qquad (1\text{-}1)$$

where f_0 is the face of the m-gon M_0 and $f(i, 1), f(i, 2), \ldots, f(i, m)$ are the faces added to $\mathcal{M}$ when M_i is circumscribed onto M_{i-1} for $i \geq 1$.

Fix an integer $m \geq 3$. Suppose $\{a_n\}_{n=0}^{\infty}$ is an increasing sequence of positive integers. We define the ordered lists $b_0 = [a_0]$ and $b_i = [a_{m(i-1)+1}, a_{m(i-1)+2}, \ldots, a_{mi}]$ for $i \leq 1$, which we refer to as bins. For all $i \geq 1$, we see that b_i has size m and b_0 has size 1. The integers in bin b_i correspond directly with the integers which we place on the faces added to $\mathcal{M}$ when M_i is circumscribed onto M_{i-1}. With the elements of our sequence partitioned into bins, we define a *legal m-gonal decomposition of*

any positive integer z. If we have

$$z = a_{\ell_t} + a_{\ell_{t-1}} + \cdots + a_{\ell_2} + a_{\ell_1}, \tag{1-2}$$

where $\ell_1 < \ell_2 < \cdots < \ell_t$ and $\{a_{\ell_j}, a_{\ell_{j+1}}\} \not\subset b_i$ for any $i \geq 0$ and $1 \leq j \leq t - 1$, then we call this a legal m-gonal decomposition of z. Namely, a legal m-gonal decomposition cannot use more than one summand from the same bin. With the generalized construction of the sequence by partitioning the members into bins rather than relying solely on the 2-dimensional circumscribed polygons, we make it a formal definition for $m \geq 1$.

Definition 1.1. Let an increasing sequence of positive integers $\{a_n\}_{n=0}^{\infty}$ be given and partition the elements into ordered lists

$$b_0 = [a_0], \quad b_k := [a_{m(k-1)+1}, a_{m(k-1)+2}, \ldots, a_{mk}] \tag{1-3}$$

for $m \geq 1$, $k \geq 1$, which we call bins. We declare a decomposition of an integer

$$z = a_{\ell_t} + a_{\ell_{t-1}} + \cdots + a_{\ell_1}, \tag{1-4}$$

where $\ell_1 < \ell_2 < \cdots < \ell_t$ and $\{a_{\ell_j}, a_{\ell_{j+1}}\} \not\subset b_i$ for any i, j, to be a *legal m-gonal decomposition*.

The following definition details the construction of the m-gonal sequence, which is the focus of this paper.

Definition 1.2. For $m \geq 1$, an increasing sequence of positive integers $\{a_n\}_{n=0}^{\infty}$ is called an *m-gonal sequence* if every a_i ($i \geq 0$) is the smallest positive integer that does not have a legal m-gonal decomposition using the elements $\{a_0, a_1, \ldots, a_{i-1}\}$.

Example 1.3. For $m = 1$, all the bins have size 1 and the 1-gonal sequence $\{a_i\}_{i=0}^{\infty}$ is defined by $a_i = 2^i$. This is equivalent to writing an integer in binary. When $m = 2$ we have bins $b_i = [a_{2i-1}, a_{2i}]$ for $i \geq 1$ and $b_0 = [a_0]$. The first few terms of the sequence are

$$\underbrace{1}_{b_0}, \ \underbrace{2, 4}_{b_1}, \ \underbrace{6, 12}_{b_2}, \ \underbrace{18, 36}_{b_3}, \ \underbrace{54, 108}_{b_4}, \ \underbrace{162, 324}_{b_5}, \ \ldots.$$

In the case where $m = 3$ the triangle (3-gonal) sequence begins with the terms

$$\underbrace{1}_{b_0}, \ \underbrace{2, 4, 6}_{b_1}, \ \underbrace{8, 16, 24}_{b_2}, \ \underbrace{32, 64, 96}_{b_3}, \ \underbrace{128, 256, 384}_{b_4}, \ \underbrace{512, 1024, 1536}_{b_5}, \ \ldots.$$

Figure 2 gives a visualization of the beginning of the triangle sequence when the integers are placed in the faces of the circumscribed triangles. Moreover, we note that the triangles used need not be equilateral.

Also one can observe that the triangle decomposition of 2015 is given by

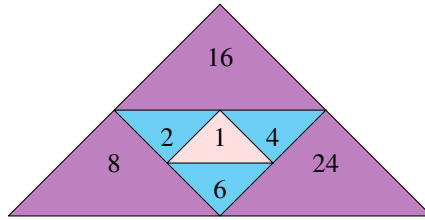$$2015 = a_{15} + a_{12} + a_6 + a_3 + a_0 = 1536 + 384 + 64 + 24 + 6 + 1. \tag{1-5}$$

Figure 2. Beginning of triangle sequence.

In Section 2 we derive the recurrence relation and explicit closed form expressions for the terms of the m-gonal sequence, which we state below.

Theorem 1.4. *Let $m \geq 1$. If $\{a_n\}_{n=0}^{\infty}$ is the m-gonal sequence, then*

$$a_n = \begin{cases} 1 & \text{if } n = 0, \\ 2n & \text{if } 1 \leq n \leq m, \\ (m+1)a_{n-m} & \text{if } n > m. \end{cases} \tag{1-6}$$

Then for $n \geq 1$, with $n = km + r$, $k \geq 0$ and $1 \leq r \leq m$,

$$a_n = 2r(m+1)^k. \tag{1-7}$$

Uniqueness of decomposition. Notice that for $m \geq 2$, the recurrence given in Theorem 1.4 is not a positive linear recurrence, as the leading coefficients of the first m terms are zero. Therefore past results on positive linear recurrences do not apply to the m-gonal sequence; however, we do still obtain unique decomposition.

Theorem 1.5 (uniqueness of decompositions). *Fix $m \geq 1$. Every positive integer can be written uniquely as a sum of distinct terms from the m-gonal sequence, where no two summands are in the same bin.*

A proof of Theorem 1.5 is given in the Appendix.

Gaussianity. Previous work with positive linear recurrence sequences proved the number of summands in the decomposition of positive integers converges to a Gaussian (see among others [Demontigny et al. 2014a; Miller and Wang 2014]). The same holds for Kentucky decompositions, despite the fact that the Kentucky sequence is not a positive linear recurrence [Catral et al. 2014], and also for the m-gonal sequences.

Theorem 1.6 (Gaussian behavior of summands). *Let the random variable Y_n denote the number of summands in the (unique) m-gonal decomposition of an integer picked at random from $[0, a_{mn+1})$ with uniform probability.[3] Normalize Y_n to*

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[3]Using the methods of [Best et al. 2016], these results can be extended to hold almost surely for a sufficiently large subinterval of $[0, a_{mn+1})$.

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$Y'_n = (Y_n - \mu_n)/\sigma_n$, where μ_n and σ_n are the mean and variance of Y_n respectively. Then

$$\mu_n = \frac{mn}{m+1} + \frac{1}{2}, \quad \sigma_n^2 = \frac{mn}{(m+1)^2} + \frac{1}{4}, \tag{1-8}$$

and Y'_n converges in distribution to the standard normal distribution as $n \to \infty$.

The proof of Theorem 1.6 is given in Section 3.

Gaps between summands. Another property of positive linear recurrence sequences that is studied is the behavior of the gaps between adjacent summands in decompositions. In many instances, it has been shown that there is exponential decay in the distribution of gaps; see [Beckwith et al. 2013; Ben-Ari and Miller 2014; Bower et al. 2015].[4] Similarly, the Kentucky sequence displays exponential decay in the distribution of gaps [Catral et al. 2014]. We obtain similar behavior again, though now there is a slight dependence on the residue of gap modulo m (if we split by residue we obtain geometric decay).

Before stating our result we first fix some notation. For the legal m-gonal decomposition

$$z = a_{\ell_k} + a_{\ell_{k-1}} + \cdots + a_{\ell_1} \quad \text{with } \ell_1 < \ell_2 < \cdots < \ell_k \tag{1-9}$$

and $z \in [0, a_{mn+1})$, we define the multiset of gaps as

$$\text{Gaps}_n(z) := \{\ell_2 - \ell_1, \ell_3 - \ell_2, \ldots, \ell_k - \ell_{k-1}\}. \tag{1-10}$$

Observe that we do not consider $\ell_1 - 0$ as a gap. However, doing so would not affect the limiting behavior. For example, notice $z = a_{15} + a_{12} + a_6 + a_3 + a_0$ contributes three gaps of length 3, and one gap of length 6.

Considering all the gaps between summands in legal m-gonal decompositions of all $z \in [0, a_{mn+1})$, we let $P_n(g)$ be the fraction of all these gaps that are of length g. That is, $P_n(g)$ is the probability of a gap of length g among legal m-gonal decompositions of $z \in [0, a_{mn+1})$.

Theorem 1.7 (average gap measure). *Let $g = m\alpha + \beta$, where $\alpha \geq 0$ and $0 \leq \beta < m$. For $P_n(g)$ as defined above, the limit $P(g) := \lim_{n \to \infty} P_n(g)$ exists, and*

$$P(g) = \begin{cases} \frac{\beta}{m(m+1)} & \text{if } \alpha = 0, \\ \frac{m+1-\beta}{(m+1)^{\alpha+1}} & \text{if } \alpha > 0. \end{cases} \tag{1-11}$$

The proof for Theorem 1.7 is given in Section 4.

Via an application of [Dorward et al. 2015, Theorem 1.1] we extract a result on individual gaps for the m-gonal case. In order to state the theorem, we need the

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[4]The proofs involve technical arguments concerning roots of polynomials associated to the recurrence; in many cases one needs to assume all the recurrence coefficients are positive.

following definitions, as were presented in the above reference, but specialized to the m-gonal case. Let $\{a_n\}$ denote the m-gonal sequence with its unique decomposition as given in Definition 1.1. Let $I_n := [0, a_{mn+1})$ for all $n > 0$ and let $\delta(x - a)$ denote the Dirac delta functional, assigning a mass of 1 to $x = a$ and 0 otherwise.

• *Spacing gap measure:* We define the spacing gap measure of a $z \in I_n$ with $k(z)$ summands as

$$\nu_{z,n}(x) := \frac{1}{k(z) - 1} \sum_{j=2}^{k(z)} \delta(x - (\ell_j - \ell_{j-1})). \qquad (1\text{-}12)$$

• *Average spacing gap measure:* Note that the total number of gaps for all $z \in I_n$ is

$$N_{\text{gaps}}(n) := \sum_{z=a_0}^{a_{mn+1}-1} (k(z) - 1). \qquad (1\text{-}13)$$

The average spacing gap measure for all $z \in I_n$ is

$$\nu_n(x) := \frac{1}{N_{\text{gaps}}(n)} \sum_{z=a_0}^{a_{mn+1}-1} \sum_{j=2}^{k(z)} \delta(x - (\ell_j - \ell_{j-1}))$$

$$= \frac{1}{N_{\text{gaps}}(n)} \sum_{z=a_0}^{a_{mn+1}-1} (k(z) - 1)\nu_{z,n}(x). \qquad (1\text{-}14)$$

Letting $P_n(g)$ denote the probability of a gap of length g among all gaps from the decompositions of all $z \in I_n$, we have

$$\nu_n(x) = \sum_{g=0}^{mn} P_n(g)\delta(x - g). \qquad (1\text{-}15)$$

• *Limiting average spacing gap measure, limiting gap probabilities:* If the limits exist, we let

$$\nu(x) = \lim_{n \to \infty} \nu_n(x), \quad P(k) = \lim_{n \to \infty} P_n(k). \qquad (1\text{-}16)$$

• *Indicator function for two gaps:* For $g_1, g_2 \geq 0$,

$$X_{j_1, j_1+g_1, j_2, j_2+g_2}(n) := \#\{z \in I_n : a_{j_1}, a_{j_1+g_1}, a_{j_2}, a_{j_2+g_2} \text{ in } z\text{'s decomposition,}$$
$$\text{but not } a_{j_1+q}, a_{j_2+p} \text{ for } 0 < q < g_1, 0 < p < g_2\}. \qquad (1\text{-}17)$$

• *Specific gap length probability:* Recall that $P_n(g)$ is the probability

$$P_n(g) := \frac{1}{N_{\text{gaps}}(n)} \sum_{i=1}^{mn+1-g} X_{i,i+g}(n). \qquad (1\text{-}18)$$

Now can now state the result of the individual gap measure for the m-gonal case.

Theorem 1.8 (individual gap measure). *For $z \in I_n$, the individual gap measures $v_{z,n}(x)$ converge almost surely in distribution to the limiting gap measure $v(x)$.*

We give a proof of Theorem 1.8 in Section 5.

Longest gap. Another interesting problem is to determine the distribution of the longest gap between summands as $n \to \infty$. The structure of the legal m-gonal decompositions allows us to easily prove the following.

Theorem 1.9 (distribution of the longest gap). *Consider the m-gonal sequence $\{a_n\}$. Then as $n \to \infty$ the mean of the longest gap between summands in legal m-gonal decompositions of integers in $[a_n, a_{n+1})$ is $m \log_2(n/2m) + O_m(1)$, and the variance is $O_m(1)$.*

The proof of Theorem 1.9 is given in Section 6 and bypasses many of the technical arguments used in [Bower et al. 2015]. There the authors had to deduce properties of somewhat general associated polynomials; the nature of the legal m-gonal decompositions here allows us to immediately convert this problem to a simple generalization of the longest run of heads problem.

2. Recurrence relations and generating functions

Let $m \geq 1$. We can use the division algorithm to observe that the integer a_{mk+r} is the r-th integer in the bin b_{k+1} for $mk + r \geq 1$. Hence $1 \leq r \leq m$ denotes the location of the integer within its bin. We let the first bin b_0 contain the element $a_0 = 1$. Then for any $k \geq 0$, we let b_{k+1} denote the set of elements of the $(k+1)$-th bin. Namely

$$\underbrace{a_0}_{b_0}, \underbrace{a_1, a_2, \ldots, a_m}_{b_1}, \underbrace{a_{m+1}, a_{m+2}, \ldots, a_{2m}}_{b_2}, \ldots, \underbrace{a_{mk+1}, a_{mk+2}, \ldots, a_{m(k+1)}}_{b_{k+1}}, \ldots. \tag{2-1}$$

We can now begin our work in describing the terms of this sequence.

The following result, which follows immediately from the definition, is used in many of the proofs in this section. We record it here for easy reference.

Definition 2.1. Let Ω_n denote the largest integer with summands from each bin $b_0, b_1, b_2, \ldots, b_n$. Then

$$\Omega_n = \sum_{i=0}^{n} a_{mi}. \tag{2-2}$$

The first result that makes use of Definition 2.1 is given below.

Lemma 2.2. *Let $m \geq 1$ and $k \geq 1$. If a_{mk+1} is the first entry in bin b_{k+1}, then $a_{mk+1} = a_{mk} + a_{m(k-1)+1}$.*

Proof. We note that since a_{mk+1} and $a_{m(k-1)+1}$ are the first numbers in the bins b_{k+1} and b_k, respectively, by (2-2) we have

$$a_{mk+1} = 1 + \Omega_k = 1 + \sum_{i=0}^{k} a_{mi}, \tag{2-3}$$

$$a_{m(k-1)+1} = 1 + \Omega_{k-1} = 1 + \sum_{i=0}^{k-1} a_{mi}. \tag{2-4}$$

Then (2-3) and (2-4) yield

$$a_{mk+1} = 1 + \Omega_k = 1 + a_{mk} + \Omega_{k-1} = a_{mk} + (1 + \Omega_{k-1}) = a_{mk} + a_{m(k-1)+1}, \tag{2-5}$$

as claimed. $\qquad\square$

We now prove a more general result.

Lemma 2.3. *If $k \geq 0$ and $1 \leq r \leq m$, then $a_{mk+r} = r a_{mk+1}$.*

Proof. First consider the bin b_1. As $b_0 = [a_0] = 1$, by construction of the m-gonal sequence it is straightforward to determine that $b_1 = [2, 4, \ldots, 2m]$ and $a_r = r a_1$ for all $1 \leq r \leq m$.

We proceed for bins b_k with $k \geq 1$ by induction on r, where $1 \leq r \leq m$. The base case $r = 1$ clearly holds.

Let $1 \leq x \leq m - 1$ and assume that for any $1 \leq r \leq x$, we have $a_{mk+r} = r a_{mk+1}$. We want to show that $a_{mk+x+1} = (x+1)a_{mk+1}$. Recall that a_{mk+x+1} is the entry in bin b_{k+1} after a_{mk+x} and by definition a_{mk+x+1} is one more than the largest integer we can create using the elements of bins $b_0, b_1, \ldots, b_k$ along with the element a_{mk+x}. Using (2-2), we have

$$a_{mk+x+1} = 1 + a_{mk+x} + \Omega_k = 1 + a_{mk+x} + \sum_{i=0}^{k} a_{mi}. \tag{2-6}$$

Since $1 + \Omega_k = a_{mk+1}$ and by the use of the induction hypothesis, (2-6) yields

$$a_{mk+x+1} = a_{mk+x} + a_{mk+1} = x a_{mk+1} + a_{mk+1} = (x+1)a_{mk+1} \tag{2-7}$$

as desired. $\qquad\square$

We now provide a closed formula for the terms of the m-gonal sequence.

Proposition 2.4. *Let $m \geq 2$, $k \geq 0$, and $1 \leq r \leq m$. Then $a_{mk+r} = 2r(m+1)^k$. For $m = 1$, we have $a_i = 2^i$.*

Proof. For the case where $m = 1$, each of our bins have size 1 and a legal decomposition has distinct summands. Thus the rule for legal decomposition is precisely a description of writing the positive integers in binary.

We will proceed by induction on k, the subscript on the bin, and r, the location of a_{mk+r} within the bin considered. The base case $k = 0$ and $r = 1$ gives the expected result, $a_{m \cdot 0 + 1} = 2(1)(m + 1)^0 = 2$. We now assume that for some $k \geq 0$ and some r, $1 \leq r \leq m$, we have

$$a_{mk+r} = 2r(m + 1)^k. \tag{2-8}$$

We need to show that the following two equations hold:

$$a_{mk+r+1} = 2(r + 1)(m + 1)^k, \tag{2-9}$$

$$a_{m(k+1)+r} = 2r(m + 1)^{k+1}. \tag{2-10}$$

Suppose $1 \leq r \leq m - 1$. To show (2-9) holds it suffices to observe that by Lemma 2.3 and our induction hypothesis we have

$$a_{mk+r+1} = (r + 1)a_{mk+1} = (r + 1) \cdot 2(1)(m + 1)^k = 2(r + 1)(m + 1)^k. \tag{2-11}$$

When $r = m$, we use Lemma 2.2 and our induction hypothesis to quickly deduce

$$a_{mk+m+1} = 2(m + 1)^{k+1}. \tag{2-12}$$

By iterating (2-11) until $r = m - 1$, we find that (2-12) holds. Then (2-10) holds by Lemma 2.3. □

The final result gives the recurrence relation stated in Theorem 1.4.

Corollary 2.5. *If $n > m$, then $a_n = (m + 1)a_{n-m}$.*

Proof. Let $n > m$ and write $n = mk + r$, where $k \geq 1$ and $1 \leq r \leq m$. By Proposition 2.4, $a_n = a_{mk+r} = 2r(m + 1)^k$ and $a_{n-m} = a_{m(k-1)+r} = 2r(m + 1)^{k-1}$. So it directly follows that $a_n = (m + 1)a_{n-m}$. □

Counting integers with exactly k summands. Koloğlu, Kopp, Miller and Wang [Koloğlu et al. 2011] introduced a very useful combinatorial perspective to attack Zeckendorf decomposition problems by partitioning the integers $z \in [F_n, F_{n+1})$ into sets based on the number of summands in their Zeckendorf decompositions. We use a similar technique to prove that the distribution of the average number of summands in the m-gonal decomposition displays Gaussian behavior.

Let $p_{n,k}$ denote the number of integers in $I_n := [0, a_{mn+1})$ whose m-gonal decomposition contains exactly k summands, where $k \geq 0$. We begin our analysis with the following result.

Proposition 2.6. *If $n, k \geq 0$, then*

$$p_{n,k} = \begin{cases} 1 & \text{if } k = 0, \\ m^k \binom{n}{k} + m^{k-1} \binom{n}{k-1} & \text{if } 1 \leq k \leq n + 1, \\ 0 & \text{if } k > n + 1. \end{cases} \tag{2-13}$$

Proof. Let $n, k \geq 0$. Observe that the unique integer in the interval $I_n = [0, a_{mn+1})$ which has zero summands is zero itself. Thus $p_{n,0} = 1$. Now if k is larger than the number of available bins, it would be impossible to have k summands, as one can draw no more than one summand per bin. Therefore $p_{n,k} = 0$ whenever $k > n + 1$.

We now show that for $1 \leq k \leq n + 1$, we have $p_{n,k} = m^k \binom{n}{k} + m^{k-1} \binom{n}{k-1}$. There are two cases to consider:

- **Case 1:** One of the k summands is chosen from b_0.

- **Case 2:** None of the k summands are chosen from b_0.

Case 1: Since one of the k summands is coming from b_0 there are $k - 1$ available summands to take from the bins $b_1, \ldots, b_n$. The number of ways to select $k - 1$ bins from n bins is $\binom{n}{k-1}$. As each of the bins $b_1, \ldots, b_n$ has exactly m elements and $|b_0| = 1$, once the $k - 1$ bins are selected, the number of ways to select an element from these bins is m^{k-1}. Thus the number of $z \in I_n$ which have exactly k summands with one summand coming from bin b_0 is $m^{k-1} \binom{n}{k-1}$.

Case 2: We choose k summands from any bin but b_0. Using a similar argument to that in Case 1, we can see that the total number of ways to select these k summands is $m^k \binom{n}{k}$.

As the two cases are disjoint, we have shown that the total number of integers in the interval I_n with exactly k summands is

$$p_{n,k} = m^k \binom{n}{k} + m^{k-1} \binom{n}{k-1} \tag{2-14}$$

as desired. $\square$

We also provide a recursive formula for the value of $p_{n,k}$, as it is used in the proof of Proposition 2.8.

Proposition 2.7. *If $0 < k < n + 1$, then $p_{n,k} = m p_{n-1,k-1} + p_{n-1,k}$.*

We omit the proof of Proposition 2.7, as it is a straightforward application of the combinatorial identity $\binom{n}{k} = \binom{n-1}{k} + \binom{n-1}{k-1}$. With the recursive formula at hand, we now create a generating function for $p_{n,k}$.

Proposition 2.8. *Let*

$$F(x, y) := \sum_{n,k \geq 0} p_{n,k} x^n y^k \tag{2-15}$$

be the generating function of the $p_{n,k}$ arising from m-gonal decompositions. Then

$$F(x, y) = \frac{1 + y}{1 - (my + 1)x}. \tag{2-16}$$

Proof. Noting that $p_{n,k} = 0$ if either $n < 0$ or $k < 0$, using explicit values of $p_{n,k}$ and the recurrence relation from Proposition 2.7, and after some straightforward algebra, we obtain

$$F(x, y) = mxyF(x, y) + xF(x, y) + 1 + y. \tag{2-17}$$

From this, (2-16) follows.   □

3. Gaussian behavior

To motivate this section's main result, we point the reader to the following experimental observations. Taking samples of 200,000 integers from the intervals $[0, 2 \cdot 4^{600})$, $[0, 2 \cdot 5^{600})$, $[0, 2 \cdot 6^{600})$ and $[0, 2 \cdot 7^{600})$, in Figure 3 we provide a histogram for the distribution of the number of summands in the m-gonal decomposition of these integers, when $m = 3, 4, 5$ and 6, respectively. Moreover, Figure 3 provides the histograms and Gaussian curves (associated to the respective value of m and n; the interval is $[0, a_{mn+1})$ so $n = 600$ in all experiments). In Table 1 we give the values of the predicted means and variances (as computed using Proposition 3.2), as well as the sample means and variances, for each of the cases considered.

From these observations it is expected that, for any $m \geq 1$, the distribution of the number of summands in the m-gonal decompositions of integers in the interval I_n will display Gaussian behavior. This is in fact the statement of Theorem 1.6. We begin by proving a technical result and follow it with the formulas for the mean and variance, which make use of some properties associated with the generating function for the $p_{n,k}$.

Proposition 3.1. *If $g_n(y)$ denotes the coefficient of x^n in $F(x, y)$, then*

$$g_n(y) = (1 + y)(my + 1)^n. \tag{3-1}$$

Proof. Using the fact that $F(x, y) = (1 + y)/(1 - mxy - x)$ we have by geometric series that

$$F(x, y) = \sum_{n=0}^{\infty} (1 + y)(my + 1)^n x^n. \tag{3-2}$$

Thus the coefficient of x^n in $F(x, y)$ is $(1 + y)(my + 1)^n$.   □

We can now use $g_n(y)$ to find the mean and variance for the number of summands for integers $z \in I_n$.

Proposition 3.2. *Let Y_n be the number of summands in the m-gonal decomposition of a randomly chosen integer in the interval I_n, where each integer has an equal probability of being chosen. Let μ_n and σ_n^2 denote the mean and variance of Y_n. Then*

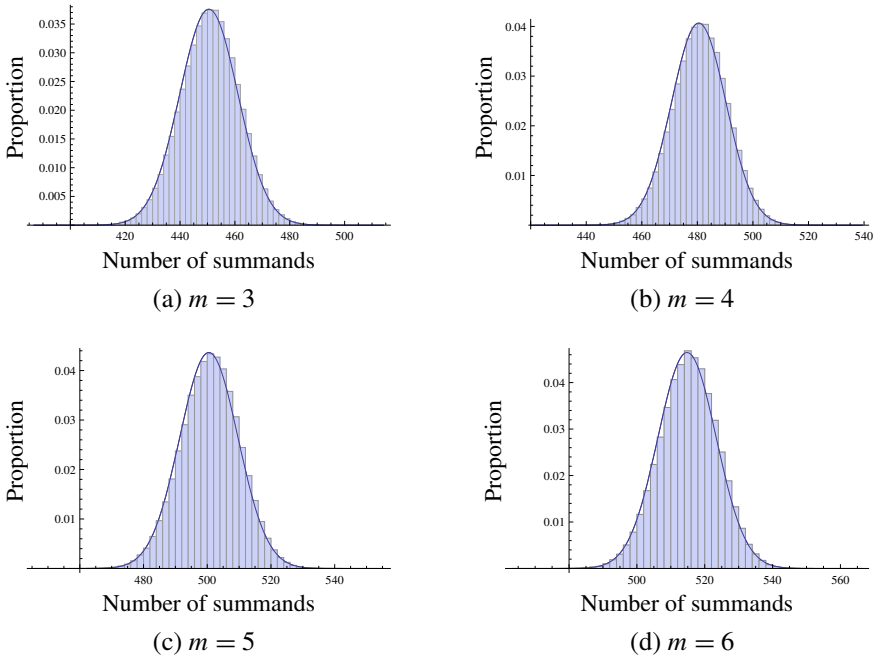$$\mu_n = \frac{nm}{m+1} + \frac{1}{2}, \quad \sigma_n^2 = \frac{nm}{(m+1)^2} + \frac{1}{4}. \tag{3-3}$$

(a) $m = 3$ (b) $m = 4$

(c) $m = 5$ (d) $m = 6$

Figure 3. Distributions for the number of summands in the m-gonal decomposition for a random sample with $n = 600$.

Figure	m	Predicted mean	Sample mean	Predicted variance	Sample variance
3(a)	3	450.50	450.49	112.75	112.34
3(b)	4	480.50	480.52	96.25	95.73
3(c)	5	500.50	450.49	83.58	83.38
3(d)	6	514.79	514.76	73.72	73.64

Table 1. Predicted means and variances versus sample means and variances for simulation from Figure 3.

Proof. By Propositions 4.7 and 4.8 in [Demontigny et al. 2014a] the mean and variance of Y_n are

$$\mu_n = \sum_{i=0}^{n} i\, P(Y_n = i) = \sum_{i=0}^{n} i\, \frac{p_{n,i}}{\sum_{k=0}^{n} p_{n,k}} = \frac{g_n'(1)}{g_n(1)}, \tag{3-4}$$

$$\sigma_n^2 = \sum_{i=0}^{n} (i - \mu_n)^2 P(Y_n = i) = \sum_{i=0}^{n} i^2 \frac{p_{n,i}}{\sum_{k=0}^{n} p_{n,k}} - \mu_n^2 = \frac{\frac{d}{dy}[y g_n'(y)]|_{y=1}}{g_n(1)} - \mu_n^2. \tag{3-5}$$

Our result follows from these formulas and the fact that $g_n(y) = (1+y)(my+1)^n$. $\qquad\square$

Normalize Y_n to $Y'_n = (Y_n - \mu_n)/\sigma_n$, where μ_n and σ_n are the mean and variance of Y_n respectively, as given in Proposition 3.2. We are now ready to prove that Y'_n converges in distribution to the standard normal distribution as $n \to \infty$.

Proof of Theorem 1.6. For convenience we set $r := t/\sigma_n$. Since

$$\sigma_n = \sqrt{\frac{nm}{(m+1)^2} + \frac{1}{4}},$$

we know that $r \to 0$ as $n \to \infty$ for any fixed value of t. Hence we will expand e^r using its power series expansion. We start with

$$M_{Y'_n}(t) = \frac{g_n(e^{t/\sigma_n})e^{-t\mu_n/\sigma_n}}{g_n(1)}. \tag{3-6}$$

Taking the logarithm of (3-6), we have

$$\log(M_{Y'_n}(t)) = \log[g_n(e^r)] - \log[g_n(1)] - \frac{t\mu_n}{\sigma_n}. \tag{3-7}$$

We proceed using Taylor expansions of the exponential and logarithmic functions to expand the following:

$$\begin{aligned}
\log[g_n(e^r)] &= \log(1+e^r) + n\log(me^r + 1) \\
&= \log\left(1 + \left(1 + r + \tfrac{1}{2}r^2\right)\right) + n\log\left(m\left(1 + r + \tfrac{1}{2}r^2\right) + 1\right) + O(r^3) \\
&= \log(2) + \tfrac{1}{2}\left(r + \tfrac{1}{2}r^2\right) - \tfrac{1}{8}\left(r + \tfrac{1}{2}r^2\right)^2 \\
&\quad + n\left(\log(m+1) + \frac{\left(mr + \tfrac{1}{2}mr^2\right)}{m+1} - \frac{\left(mr + \tfrac{1}{2}mr^2\right)^2}{2(m+1)^2}\right) + O(r^3) \\
&= \log(2(m+1)^n) + \tfrac{1}{2}r + \tfrac{1}{8}r^2 + \frac{nmr}{m+1} + \frac{nmr^2}{2(m+1)^2} + O(r^3). \tag{3-8}
\end{aligned}$$

From Proposition 3.1 we have $g_n(1) = 2(m+1)^n$, hence

$$\log[2(m+1)^n] = \log[g_n(1)]. \tag{3-9}$$

Substituting (3-8) and (3-9) and the values

$$\mu_n = \frac{nm}{m+1} + \frac{1}{2} \quad \text{and} \quad \sigma_n = \sqrt{\frac{nm}{(m+1)^2} + \frac{1}{4}}$$

into (3-7) yields

$$\log(M_{Y'_n}(t)) = \tfrac{1}{2}r + \tfrac{1}{8}r^2 + \frac{nmr}{m+1} + \frac{nmr^2}{2(m+1)^2} - \frac{t\left(\frac{nm}{m+1} + \frac{1}{2}\right)}{\sqrt{\frac{nm}{(m+1)^2} + \frac{1}{4}}} + O(r^3). \tag{3-10}$$

After some straightforward algebra we arrive at

$$\log(M_{Y_n'}(t)) = \tfrac{1}{2}t^2 + o(1); \tag{3-11}$$

the moment-generating proof of the central limit theorem now yields that the distribution converges to that of the standard normal distribution as $n \to \infty$. $\square$

4. Average gap measure

We now turn our attention to our final result in which we determine the behavior of gaps between summands. We begin with some preliminary notation in order to make our approach precise. For a positive integer $z \in I_n = [0, a_{mn+1})$ with m-gonal decomposition

$$z = a_{\ell_t} + a_{\ell_{t-1}} + \cdots + a_{\ell_1}, \tag{4-1}$$

where $\ell_1 < \ell_2 < \cdots < \ell_t$, we define the multiset of gaps of z as

$$\text{Gaps}_n(z) := \{\ell_2 - \ell_1, \ell_3 - \ell_2, \ldots, \ell_t - \ell_{t-1}\}. \tag{4-2}$$

Our result will average over all $z \in [0, a_{mn+1})$ since we are interested in the average gap measure arising from m-gonal decompositions.

We follow the methods of [Beckwith et al. 2013; Bower et al. 2015]. In order to have a gap of length exactly g in the decomposition of z, there must be some index i such that a_i and a_{i+g} occur in z's decomposition, but a_j does not for any j between i and $i + g$. Thus for each i we count how many z have a_i and a_{i+g} but not a_j for $i < j < i + g$; summing this count over i gives the number of occurrences of a gap of length g among all the decompositions of z in our interval of interest. We want to compute the fraction of the gaps (of length g) arising from the decompositions of all $z \in I_n$. This probability is given by

$$P_n(g) := \frac{1}{(\mu_n - 1)a_{mn+1}} \sum_{z=0}^{a_{mn+1}-1} \sum_{i=0}^{mn+1-g} X_{i,g}(z), \tag{4-3}$$

where $X_{i,g}(z)$ is the indicator function.[5]

We are now ready to prove the result on the exponential decay in the distribution of gaps. The arguments in the proof of our main result (Theorem 1.7) are quite straightforward, but a bit tedious. To simplify our arguments, we write the gap length g as $m\alpha + \beta$, where $\alpha \geq 0$ and $0 \leq \beta < m$.

---

[5]For $1 \leq i, g \leq mn$ $X_{i,g}(z)$ denotes whether the decomposition of z has a gap of length g beginning at index i. That is, for $z = a_{\ell_t} + a_{\ell_{t-1}} + \cdots + a_{\ell_1}$,

$$X_{i,g}(z) = \begin{cases} 1 & \text{if } \exists \, j, 1 \leq j \leq t \text{ with } i = \ell_j \text{ and } i + g = \ell_{j+1}, \\ 0 & \text{otherwise.} \end{cases}$$

Proof of Theorem 1.7. Let $g = m\alpha + \beta$, where $\alpha \geq 0$ and $0 \leq \beta < m$. We proceed by considering the following two cases:

- **Case 1:** $\alpha = 0$.

- **Case 2:** $\alpha > 0$.

Case 1: Let $\alpha = 0$. Hence $g = \beta$, where $0 < \beta < m$, and so our gap is less than the size of each bin b_i for $i > 0$ ($\alpha = 0$ and $\beta = 0$ would give us a gap of length 0 which is not m-gonal legal). We first consider gaps of length $g = \beta$ beginning at index 0. If $a_i = a_0$ then the only way to have a gap of length $g = \beta$ is if $a_{i+1} = a_\beta$. Now we are counting integers with m-gonal decompositions of the form $a_{\ell_t} + \cdots + a_{\ell_3} + a_\beta + a_0$. The number of $z \in I_n$ with summands $a_{\ell_3}, \ldots, a_{\ell_t}$ coming from bins $b_2, b_3, \ldots, b_n$ is $(m+1)^{n-1}$.

If $a_i = a_{mk+r}$, $k \geq 0$ and $1 \leq r \leq m$, then $a_{i+g} = a_{mk+r+\beta}$. Notice that the case $r + \beta \leq m$ cannot occur as this would force $\{a_i, a_{i+g}\} \subset b_{k+1}$, which leads to a decomposition that is not m-gonal legal, as only one summand can be taken per bin. Now if $r + \beta > m$, then $a_{i+g} \in b_{k+2}$.

Notice that in this case $a_i = a_{mk+r}$ is one of the largest β entries in bin b_{k+1} and thus there are β many choices for r, namely $r \in \{m-\beta+1, m-\beta+2, \ldots, m\}$.

Now we need to count the number of $z \in I_n \setminus \{0\} = [1, a_{mn+1})$ which have summands $a_i \in b_{k+1}$ and $a_{i+g} \in b_{k+2}$. We must have $0 \leq k \leq n - 2$.

As we have already used bins b_{k+1} and b_{k+2}, it follows from a straightforward combinatorial counting argument that the total number of integers $z \in I_n$ that can be created with summands $a_i \in b_{k+1}$ and $a_{i+g} \in b_{k+2}$ (with no summands in between) is given by

$$2\beta(m+1)^{n-2}, \tag{4-4}$$

where the factor of β comes from the β possible choices of a_i within the bin b_{k+1}. As we can vary k from 0 to $n - 2$, we find that the total number of integers $z \in I_n$ which contribute a gap between a_i ($i \neq 0$) and a_{i+g} (assuming that $r + \beta > m$) is given by

$$2(n-1)\beta(m+1)^{n-2}. \tag{4-5}$$

Observe that (4-5) does not account for the case when $i = 0$, which adds an extra factor of $(m+1)^{n-1}$. Therefore

$$\sum_{z=0}^{a_{mn+1}-1} \sum_{i=0}^{mn+1-g} X_{i,g}(z) = 2\beta(n-1)(m+1)^{n-2} + (m+1)^{n-1}. \tag{4-6}$$

Using Proposition 3.2 and Proposition 2.4 we have

$$(\mu_n - 1)a_{mn+1} = \left(\frac{mn}{m+1} - \frac{1}{2}\right)(2(m+1)^n) = (m+1)^{n-1}(2mn - (m+1)). \tag{4-7}$$

By (4-6) and (4-7), for $g = \beta$, with $0 \leq \beta < m$, we have

$$P_n(g) = \frac{2\beta(n-1)}{(m+1)(2mn - m - 1)} + \frac{1}{2mn - m - 1}. \tag{4-8}$$

Now recall that $P(g) = \lim_{n \to \infty} P_n(g)$, so by letting $n \to \infty$ in (4-8) we have

$$P(g) = \frac{\beta}{m(m+1)} \tag{4-9}$$

whenever $g = \beta$ and $0 \leq \beta < m$. This completes Case 1.

Case 2: Let $g = m\alpha + \beta$, where $\alpha \geq 1$ and $0 \leq \beta < m$. First consider when $a_i = a_0$. If $\beta = 0$, then $a_{i+g} = a_{m\alpha} \in b_\alpha$. Otherwise, when $0 < \beta < m$, we have $a_{i+g} = a_{m\alpha+\beta} \in b_{\alpha+1}$. In the case of the former, the number of $z \in I_n$ with summands coming from bins $b_{\alpha+1}, b_{\alpha+2}, \ldots, b_n$ is $(m+1)^{n-\alpha}$. In the latter case, the number of $z \in I_n$ with summands coming from bins $b_{\alpha+2}, b_{\alpha+3}, \ldots, b_n$ is $(m+1)^{n-(\alpha+1)}$.

Now if $a_i = a_{mk+r}$, with $k \geq 0$ and $1 \leq r \leq m$, then $a_{i+g} = a_{m(k+\alpha)+r+\beta}$. Hence $a_i \in b_{k+1}$ and $a_{i+g} \in b_{k+\alpha+1}$ whenever $1 \leq r + \beta \leq m$, or $a_{i+g} \in b_{k+\alpha+2}$ whenever $m < r + \beta < 2m$. Hence we consider the following subcases:

- **Subcase 1:** Let $1 \leq r + \beta \leq m$.
- **Subcase 2:** Let $m < r + \beta < 2m$.

Subcase 1: Let $1 \leq r + \beta \leq m$. In this case $a_i \in b_{k+1}$ and $a_{i+g} \in b_{k+\alpha+1}$. Notice that in this case a_i must be one of the smallest $m - \beta$ entries in bin b_{k+1}. Namely $r = 1, 2, \ldots, m - \beta$.

Now we need to count the number of $z \in I_n$ which have summands $a_i \in b_{k+1}$ and $a_{i+g} \in b_{k+\alpha+1}$ (and no summands in between). For the decomposition to only have summands from bins $b_0, \ldots, b_{k+1}, b_{k+\alpha+1}, \ldots, b_n$, we must have $0 \leq k \leq n-(\alpha+1)$.

In order to have a gap created by $a_i \in b_{k+1}$ and $a_{i+g} \in b_{k+\alpha+1}$, there must be no summands taken from b_j, where $k+1 < j < k+\alpha+1$. Again using a straightforward combinatorial counting argument, the total number of integers $z \in I_n$ which have summands $a_i \in b_{k+1}$ and $a_{i+g} \in b_{k+\alpha+1}$ (with no summands in between) is given by

$$2(m - \beta)(m+1)^{n-\alpha-1}, \tag{4-10}$$

where the factor of $m - \beta$ comes from the $m - \beta$ possible choices of a_i within the bin b_{k+1}.

As we can vary k from 0 to $n - (\alpha + 1)$ we find that the total number of integers $z \in I_n$ which contribute a gap between a_i $(i \neq 0)$ and a_{i+g} in this case is

$$2(n - \alpha)(m - \beta)(m+1)^{n-\alpha-1}. \tag{4-11}$$

Subcase 2: Let $m < r + \beta < 2m$. In this case $a_i \in b_{k+1}$ and $a_{i+g} \in b_{k+\alpha+2}$. Notice that in this case a_i can be any of the largest β entries in bin b_{k+1} so $a_i \in b_{k+1}$ and $a_{i+g} \in b_{k+\alpha+2}$. Namely $r = m+1-\beta,\ m+2-\beta,\ \ldots,\ m$.

Using the same reasoning as in Subcase 1, we determine that the total number of integers meeting the conditions is

$$2\beta(n - \alpha - 1)(m + 1)^{n-\alpha-2}. \qquad (4\text{-}12)$$

This completes Subcase 2.

We still need to account for the number of integers $z \in I_n$ which contribute a gap of length $g = m\alpha + \beta$ ($\alpha \geq 1$ and $0 \leq \beta < m$) beginning at index a_0. Recall we previously computed this quantity to be $(m + 1)^{n-\alpha-1}$ when $\beta > 0$ and the quantity is $(m + 1)^{n-\alpha}$ when $\beta = 0$.

Therefore we need to sum the values of (4-11) and (4-12) along with $(m+1)^{n-\alpha-1}$ when $\beta > 0$ to get

$$\sum_{z=0}^{a_{mn+1}-1} \sum_{i=0}^{mn+1-g} X_{i,g}(z) = 2(n-\alpha)(m-\beta)(m+1)^{n-\alpha-1}$$
$$+2\beta(n-\alpha-1)(m+1)^{n-\alpha-2}+(m+1)^{n-\alpha-1}. \quad (4\text{-}13)$$

By (4-13) and (4-7), for $g = m\alpha + \beta$, with $\alpha \geq 1$ and $0 < \beta < m$, we have

$$P_n(g) = \frac{2(n - \alpha)(m - \beta)}{(m + 1)^\alpha (2mn - m - 1)} + \frac{2\beta(n - \alpha - 1)}{(m + 1)^{\alpha+1}(2mn - m - 1)}$$
$$+ \frac{1}{(m + 1)^\alpha (2mn - m - 1)}. \quad (4\text{-}14)$$

Now recall that $P(g) = \lim_{n\to\infty} P_n(g)$, so by letting $n \to \infty$ in (4-14) we have

$$P(g) = \frac{2(m - \beta)}{(m + 1)^\alpha (2m)} + \frac{2\beta}{(m + 1)^{\alpha+1}(2m)} = \frac{m + 1 - \beta}{(m + 1)^{\alpha+1}} \qquad (4\text{-}15)$$

whenever $g = m\alpha + \beta$, $\alpha \geq 1$ and $0 < \beta < m$.

Now for $\beta = 0$ we do not need to consider when $\alpha = 0$, as this would give us a gap $g = 0$. Also, as $1 \leq r \leq m$, we only meet the conditions of Subcase 1. Thus for $\alpha > 0$ and $\beta = 0$ we need to sum the values of (4-11) along with $(m + 1)^{n-\alpha}$ to get

$$\sum_{z=0}^{a_{mn+1}-1} \sum_{i=0}^{mn+1-g} X_{i,g}(z) = 2(n-\alpha)m(m + 1)^{n-\alpha-1} + (m + 1)^{n-\alpha}. \qquad (4\text{-}16)$$

By (4-16) and (4-7), for $g = m\alpha$, with $\alpha \geq 1$, we have

$$P_n(g) = \frac{2(n - \alpha)m}{(m + 1)^\alpha (2mn - m - 1)} + \frac{1}{(m + 1)^{\alpha-1}(2mn - m - 1)}. \qquad (4\text{-}17)$$

Now recall that $P(g) = \lim_{n\to\infty} P_n(g)$, so by letting $n \to \infty$ in (4-17) we have

$$P(g) = \frac{1}{(m + 1)^\alpha} = \frac{m + 1}{(m + 1)^{\alpha+1}} \qquad (4\text{-}18)$$

whenever $g = m\alpha + \beta$, $\alpha \geq 1$ and $\beta = 0$. $\qquad \square$

5. Individual gap measure

In this section, we prove Theorem 1.8 by checking that the conditions given in [Dorward et al. 2015, Theorem 1.1] are satisfied in the m-gonal case. We restate this theorem below for ease of reference.

Theorem 5.1 [Dorward et al. 2015]. *For $z \in I_n$, the individual gap measures $v_{z,n}(x)$ converge almost surely in distribution to the average gap measure $v(x)$ if the following hold*:

(1) *The number of summands for decompositions of $z \in I_n$ converges to a Gaussian with mean $\mu_n = c_{\mathrm{mean}} n + O(1)$ and variance $\sigma_n^2 = c_{\mathrm{var}} n + O(1)$ for constants $c_{\mathrm{mean}}, c_{\mathrm{var}} > 0$, and $k(z) \ll n$ for all $z \in I_n$.*

(2) *We have the following, with $\lim_{n \to \infty} \sum_{g_1, g_2} \mathrm{error}(n, g_1, g_2) = 0$:*

$$\frac{2}{|I_n| \mu_n^2} \sum_{j_1 < j_2} X_{j_1, j_1 + g_1, j_2, j_2 + g_2}(n) = P(g_1) P(g_2) + \mathrm{error}(n, g_1, g_2). \qquad (5\text{-}1)$$

(3) *The limits in* (1-16) *exist.*

The above theorem is more general than we need, and in our particular case our interval of interest is $I_n = [0, a_{mn+1})$. Now observe that Proposition 3.2 and Theorem 1.6 ensure that the first criterion is met, and $k(z)$ is clearly at most $mn + 1$ and thus $k(z) \ll n$. In addition, the exponential decay seen in Theorem 1.7 shows that condition (3) is met. It remains to show that condition (2) of Theorem 5.1 holds.

Proposition 5.2. *We have*

$$\frac{2}{|I_n| \mu_n^2} \sum_{j_1 < j_2} X_{j_1, j_1 + g_1, j_2, j_2 + g_2}(n) = P(g_1) P(g_2) + \mathrm{error}(n, g_1, g_2) \qquad (5\text{-}2)$$

and the sum of the error over all pairs (g_1, g_2) goes to zero as $n \to \infty$.

Proof. Let

$$g_1 = \alpha_1 m + \beta_1, \quad g_2 = \alpha_2 m + \beta_2, \quad j_1 = k_1 m + r_1, \quad j_2 = k_2 m + r_2,$$

where $0 \le \beta_1, \beta_2 < m$, $1 \le k_1, k_2 \le m$ and $k_1 < k_2$. Thus a_{j_1} and a_{j_2} are in bins b_{k_1+1} and b_{k_2+1} respectively. There are a number of cases to consider when determining $\sum_{j_1 < j_2} X_{j_1, j_1 + g_1, j_2, j_2 + g_2}(n)$, depending on whether or not $\alpha_1 = 0$ or $\alpha_2 = 0$. We include only the case when both $\alpha_1, \alpha_2 \ge 1$, as the other cases are similar.

There are several cases to consider. We will first consider the four cases which contribute to the main term and then bound the remaining cases. In all of the cases contributing to the main term we have $j_1 + g_1 \ne j_2$ and $j_1 \ne 0$ and thus we will suppose these conditions hold below.

Case 1: Let $1 \leq r_1 + \beta_1 \leq m$ and $1 \leq r_2 + \beta_2 \leq m$. First we determine the possible values of k_1 and k_2. In this case, the gap from g_1 spans $\alpha_1 + 1$ bins and the gap from g_2 spans $\alpha_2 + 1$ bins. Thus $0 \leq k_1 \leq n - \alpha_1 - \alpha_2 - 3$ and $k_1 + \alpha_1 + 2 \leq k_2 \leq n - \alpha_2 - 1$ and the number of choices for k_1 and k_2 is $\frac{1}{2}n^2 + O(n)$. Because of the restrictions of $1 \leq r_1 + \beta_1 \leq m$ and $1 \leq r_2 + \beta_2 \leq m$, within each bin there are $m - \beta_1$ choices for where to place a_{j_1} and $m - \beta_2$ choices for where to place a_{j_2}. Lastly, we determine the number of ways to choose the remaining elements for the decomposition. Because we are spanning $\alpha_1 + 1$ bins for the gap from g_1 and $\alpha_2 + 1$ bins for the gap from g_2, using a straightforward combinatorial counting argument, the number of integers that can be made using what remains is $2(m+1)^{n-\alpha_1-\alpha_2-2}$. Thus the total number of integers that can be made in this case is

$$2(m - \beta_1)(m - \beta_2)(m + 1)^{n-\alpha_1-\alpha_2-2}(n^2/2 + O(n))$$
$$= (m - \beta_1)(m - \beta_2)(m + 1)^{n-\alpha_1-\alpha_2-2}(n^2 + O(n)). \quad (5\text{-}3)$$

Through similar arguments we can obtain the remaining three cases that contribute to the main term.

Case 2: Let $1 \leq r_1 + \beta_1 \leq m$ and $m < r_2 + \beta_2 < 2m$. Then the number of integers that can be made in this case is

$$(m - \beta_1)\beta_2(m + 1)^{n-\alpha_1-\alpha_2-3}(n^2 + O(n)). \quad (5\text{-}4)$$

Case 3: Let $m < r_1 + \beta_1 < 2m$ and $1 \leq r_2 + \beta_2 \leq m$. Then the number of integers that can be made in this case is

$$(m - \beta_2)\beta_1(m + 1)^{n-\alpha_1-\alpha_2-3}(n^2 + O(n)). \quad (5\text{-}5)$$

Case 4: Let $m < r_1 + \beta_1 < 2m$ and $m < r_2 + \beta_2 < 2m$. Then the number of integers that can be made in this case is

$$\beta_1\beta_2(m + 1)^{n-\alpha_1-\alpha_2-4}(n^2 + O(n)). \quad (5\text{-}6)$$

The remaining cases occur when $j_1 + g_1 = j_2$ or $j_1 = 0$. In these cases, the number of choices for k_1 and k_2 is on the order of n instead of n^2 and thus the number of integers that can be made in these cases is $O(n(m + 1)^{n-\alpha_1-\alpha_2})$. Combining all cases, we have

$$\sum_{j_1 < j_2} X_{j_1, j_1+g_1, j_2, j_2+g_2}(n) = n^2(m-\beta_1)(m-\beta_2)(m+1)^{n-\alpha_1-\alpha_2-2}$$
$$+ n^2(m-\beta_1)\beta_2(m+1)^{n-\alpha_1-\alpha_2-3}$$
$$+ n^2(m-\beta_2)\beta_1(m+1)^{n-\alpha_1-\alpha_2-3}$$
$$+ n^2\beta_1\beta_2(m+1)^{n-\alpha_1-\alpha_2-4}$$
$$+ O(n(m+1)^{n-\alpha_1-\alpha_2}). \quad (5\text{-}7)$$

Next, recall that by Proposition 3.2 we have $\mu_n = nm/(m+1) + \frac{1}{2}$. In addition, $|I_n| = a_{mn+1} = 2(m+1)^n$. Thus in our case we have

$$\frac{2}{|I_n|\mu_n^2} \sum_{j_1 < j_2} X_{j_1, j_1+g_1, j_2, j_2+g_2}(n)$$

$$= \frac{1}{(m+1)^n \left(\frac{nm}{m+1} + \frac{1}{2}\right)^2} \sum_{j_1 < j_2} X_{j_1, j_1+g_1, j_2, j_2+g_2}(n)$$

$$= \left(\frac{1}{(m+1)^n \left(\frac{nm}{m+1} + \frac{1}{2}\right)^2}\right)\left(n^2(m-\beta_1)(m-\beta_2)(m+1)^{n-\alpha_1-\alpha_2-2}\right.$$

$$+ n^2(m-\beta_1)\beta_2(m+1)^{n-\alpha_1-\alpha_2-3} + n^2(m-\beta_2)\beta_1(m+1)^{n-\alpha_1-\alpha_2-3}$$

$$\left. + n^2\beta_1\beta_2(m+1)^{n-\alpha_1-\alpha_2-4} + O(n(m+1)^{n-\alpha_1-\alpha_2})\right)$$

$$= \left(\frac{1}{(m+1)^{\alpha_1+\alpha_2+2}m^2n^2 + O(n)}\right)\left(n^2(m-\beta_1)(m-\beta_2)(m+1)^2\right.$$

$$\left. + n^2(m-\beta_1)\beta_2(m+1) + n^2(m-\beta_2)\beta_1(m+1) + n^2\beta_1\beta_2\right)$$

$$+ O\left(\frac{1}{n(m+1)^{\alpha_1+\alpha_2}}\right). \tag{5-8}$$

Taking the limit as $n \to \infty$ and rearranging we obtain

$$\frac{m^2(m+1-\beta_1)(m+1-\beta_2)}{m^2(m+1)^{\alpha_1+\alpha_2+2}} = P(g_1)P(g_2). \tag{5-9}$$

The fact that the error term decays exponentially in g_1 and g_2 ensures that the error summed over all g_1 and g_2 goes to zero. □

6. Longest gap

Using the techniques introduced by Bower, Insoft, Li, Miller and Tosteson [Bower et al. 2015], we can determine the mean and variance of the distribution of the longest gap between summands in the decomposition of integers in $[a_n, a_{n+1})$ as $n \to \infty$. There are no obstructions to using those methods; however, there are some bookkeeping issues due to the nature of our legal m-gonal decomposition. Specifically, we have to worry a little about the residue of the longest gap modulo m. This is a minor issue, as with probability 1 the longest gap is much larger than m and thus we will not have two items in the same bin. Rather than going through the technical argument, we instead give a very short proof that captures the correct main term of the mean of the longest gap, which grows linearly with $\log n$; our error is at the level of the constant term for the mean. We are able to handle the variance similarly, and compute that up to an error it is $O_m(1)$.

Proof of Theorem 1.9. The proof follows immediately from results on the longest run of heads in tosses of a fair coin; we sketch the details below. If a coin has probability p of heads and $q = 1 - p$ of tails, the expected longest run of heads and its variance in n tosses is

$$\mu_n = \log_{1/p}(nq) - \frac{\gamma}{\log p} - \tfrac{1}{2} + r_1(n) + \epsilon_1(n), \quad \sigma_n^2 = \frac{\pi^2}{6\log^2 p} + \tfrac{1}{12} + r_2(n) + \epsilon_2(n);$$

(6-1)

here γ is Euler's constant, the $r_i(n)$ are at most .000016, and the $\epsilon_i(n)$ tend to zero as $n \to \infty$. Note the variance is bounded independently of n (by essentially 3.5); see [Schilling 1990] for a proof.

Note that for legal m-gonal decompositions we either have an element in a bin, or we do not. As all decompositions are equally likely, we see that these expansions are equivalent to flipping a coin with probability $\tfrac{1}{2}$ for each bin, and choosing exactly one of the m possible summands in that bin if we have a tail. As the probability that the longest gap is at the very beginning or very end of a sequence of coin tosses is negligible, we can ignore the fact that the first bin has size 1 and that we may only use part of the last bin if $n + 1$ is not a multiple of m. Thus gaps between bins used in a decomposition correspond to strings of consecutive heads.

As our integers lie in $[a_n, a_{n+1})$, we have $\lfloor n/m \rfloor + O(1) = n/m + O(1)$ bins (again, we ignore the presence or absence of the initial bin of length 1 or a partial bin at the end). We now invoke the results on the length of the longest run for tosses of a fair coin. For us, this translates not to a result on the longest gap between summands, but to a result on the longest number of *bins* between summands. It is trivial to pass from this to our desired result, as all we must do is multiply by m (the error will be at most $O(m)$ coming from the location of where the summands are in the two bins). This completes the proof of our claim on the mean; the variance follows similarly. $\qquad\square$

Appendix: Proof of Theorem 1.5

Our proof is constructive. We build our sequence by only adjoining terms that ensure that we can *uniquely* decompose a number while never using more than one summand from the same bin. For a fixed $m \geq 1$ the sequence begins

$$\underbrace{1}_{b_0}, \underbrace{2, 4, 6, \ldots, 2m}_{b_1}, \underbrace{2(m+1), 4(m+1), 6(m+1), \ldots, 2m(m+1)}_{b_2}, \ldots. \quad \text{(A-1)}$$

Note we would not adjoin 7 because then there would be two legal m-gonal decompositions for 7, one using $7 = 7$ and the other being $7 = 6 + 1$. The next number in the sequence must be the smallest integer which cannot be legally decomposed using the current terms of the sequence.

We can now proceed with proof by induction. The integers $1, 2, 3, \ldots, 2m$ have unique decompositions as they are either in the sequence or are the sum of an even number from bin b_1 plus the 1 from bin b_0. The sequence continues

$$\ldots, \underbrace{a_{m(n-2)+1}, a_{m(n-2)+2}, \ldots, a_{m(n-1)}}_{b_{n-1}}, \underbrace{a_{m(n-1)+1}, \ldots, a_{mn}}_{b_n}, \underbrace{a_{mn+1}, \ldots, a_{m(n+1)}}_{b_{n+1}}, \ldots. \tag{A-2}$$

By induction we assume that there exists a unique decomposition for all integers $z \leq a_{mn} + \Omega_{n-1}$, where Ω_{n-1} is the maximum integer that can be legally decomposed using terms in the set $\{a_0, a_1, a_2, a_3, \ldots, a_{m(n-1)}\}$.

By construction we have

$$\Omega_n = a_{mn} + \Omega_{n-1} = a_{mn} + a_{m(n-1)+1} - 1.$$

Let x be the maximum integer that can be legally decomposed using terms in the set $\{a_1, a_2, a_3, \ldots, a_{m(n-1)}\}$. Note $x = a_{m(n-1)+1} - 1$, as this is why we include $a_{m(n-1)+1}$ in the sequence.

Claim. *We have $a_{mn+1} = \Omega_n + 1$ and this decomposition is unique.*

By induction we know that Ω_n is the largest value that we can legally decompose using only terms in $\{a_0, a_1, a_2, \ldots, a_{mn}\}$. Hence we choose $\Omega_n + 1$ as a_{mn+1} and $\Omega_n + 1$ has a unique decomposition.

Claim. *Every $N \in [\Omega_n+1, \Omega_n+1+x] = [a_{mn+1}, a_{mn+1}+x]$ has a unique decomposition.*

We can legally and uniquely decompose the integers $1, 2, 3, \ldots, x$ using elements in the set $\{a_0, a_1, a_2, \ldots, a_{m(n-1)}\}$. Adding a_{mn+1} to the decomposition of any of these integers would still yield a legal m-gonal decomposition since a_{mn+1} is not in any of the bins $b_0, b_1, b_2, \ldots, b_{n-1}$. The uniqueness of these decompositions follows from the fact that if a_{mn+1} is not included as a summand, then the decomposition does not yield a number big enough to exceed $\Omega_n + 1$.

Claim. *We have $a_{mn+2} = \Omega_n + 1 + x + 1 = a_{mn+1} + x + 1$ and this decomposition is unique.*

By construction the largest integer that can be legally decomposed using terms $\{a_0, a_1, a_2, \ldots, a_{mn+1}\}$ is $\Omega_n + 1 + x$.

Claim. *Every $N \in [a_{mn+2}, a_{mn+2} + x]$ has a unique decomposition.*

First note that the decomposition exists, as we can legally and uniquely construct $a_{mn+2} + v$, where $0 \leq v \leq x$. For uniqueness, we note that if we do not use a_{mn+2}, then the summation would be too small.

Claim. *We know $a_{mn+2} + x$ is the largest integer that can be legally decomposed using terms $\{a_0, a_1, a_2, \ldots, a_{mn+2}\}$.*

This follows from construction. $\qquad\square$

Acknowledgments

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